Math 461: Probability Theory

Midterm 2, Spring 2022



Date: April 13, 2022



NAME:____

NetID:_____

READ THE FOLLOWING INFORMATION.

- This is a 50-minute.
- This exam contains 8 pages (including this cover page) and 6 questions. Total of points is 60.
- Books, notes, and other aids are not allowed. Using calculator is okay.
- Show all steps to earn full credit.
- Do not unstaple pages. Loose pages will be ignored.

Question	Points	Score
1	10	
2	12	
3	10	
4	10	
5	8	
6	10	
Total:	60	

Midterm 2

- 1. (10 points) Circle True or False. Do not justify your answer.
 - (a) True False

If *X* has normal distribution with mean 7 and variance 10, then $Z = \frac{1}{10}(X - 7)$ has standard normal distribution.

$$X \sim N(\mu, \sigma^{2}), \quad M = 7, \quad G = \sqrt{10}$$

$$Z = \frac{X - \mu}{\sigma} = \frac{X - 7}{\sqrt{10}} \sim N(0, 1)$$

(b) True False If *F* is a CDF of a random variable of *X*, then $(F(x))^2$ is also a CDF of some other random variable.

If
$$X, Y$$
 are independent RVs with the
same $F(X)$, then the CDF of
 $\max \{X, Y\}$ is $F(X)^2$.

(c) True False The joint probability mass function of *X* and *Y* is given by

$$p(1,1) = p(2,1) = \frac{1}{8}, \quad p(1,2) = \frac{1}{4}, \quad p(2,2) = \frac{1}{2}.$$

Then, $\mathbb{E}[XY] = \frac{23}{8}$.

$$E[XY] = 1 \cdot 1 \cdot p(1,1) + 2 \cdot 1 \cdot p(2,1) + 1 \cdot 2 \cdot p(1,2) + 2 \cdot 2 \cdot p(2,2) = \frac{1}{8} (1+2+4+16) = \frac{23}{8}$$

If
$$f(x, y)$$
 is the joint density of random variables *X* and *Y* given by

$$f(x,y) = \begin{cases} 1, & 0 < x < 1, 0 < y < 2x \\ 0, & \text{otherwise,} \end{cases}$$

then *X* and *Y* are independent.

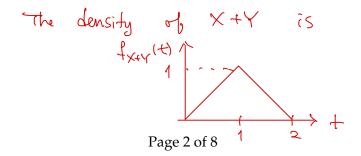
$$f_{X}(x) = \int_{0}^{1} 1 \, dx = 2x$$

$$f_{Y}(y) = \int_{1}^{1} 1 \, dx = 1 - \frac{y}{2}$$

$$f_{Y}(y) = \int_{1}^{1} y_{12} \, dx = 1 - \frac{y}{2}$$

(e) True False

If *X*, *Y* are independent uniform random variables over (0, 1), then X + Y is also a uniform random variable.



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- 2. (12 points) Give definitions or explanations of the following.
 - (a) Write down the definition of Jointly continuous random variables *X* and *Y*.

There exists
$$f(x,y) \gg 0$$
 s.t. for all $A \subseteq \mathbb{R}^{n}$
 $\mathbb{P}((X,Y) \in A) = \iint_{A} f(x,y) dxdy$.

(b) Let X, Y be jointly continuous random variables with joint density f(x, y). What is the marginal densities of X and Y?

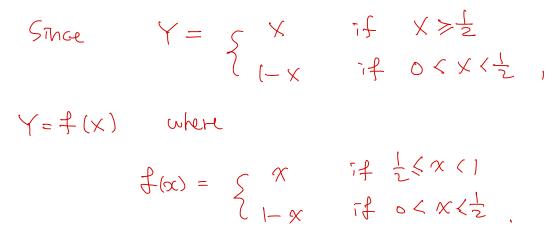
$$f_{x}(x) = \int_{R} f(x,y) \, dy$$
$$f_{y}(y) = \int_{R} f(x,y) \, dx$$

- (c) Let X, Y be jointly continuous random variables with joint density f(x, y). What is the conditional probability density functions of X given Y = y?
 - If $f_{Y}(y) > 0$, then $f_{X|Y}(x|y) = \frac{f(x,y)}{f_{Y}(y)}$
- (d) Explain Normal approximation to binomial.

If
$$X \sim Bin(n,p)$$
 $n : large, then
 $\frac{X - np}{\sqrt{np(F-p)}}$ is approximately standard Normal.$

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- 3. A stick of length 1 is broken into two pieces at *X*, where *X* is a uniform random variable over (0, 1).
 - (a) (5 points) Let *Y* be the length of the longer piece. Find a function $f : (0,1) \to \mathbb{R}$ such that Y = f(X).



(b) (5 points) Find $\mathbb{E}[Y]$.

$$E[Y] = E[f(X)] = \int_{0}^{1} f(x) dx$$
$$= \int_{0}^{\frac{1}{2}} (1-x) dx + \int_{\frac{1}{2}}^{1} x dx$$
$$= \frac{3}{4}.$$

- 4. Let *X*, *Y* be independent exponential random variables with parameters λ_1 and λ_2 respectively.
 - (a) (5 points) Is $Z := \min\{X, Y\}$ an exponential random variable? Justify your answer.

Yes.

$$F_{z}(t) = \mathbb{P}(z \leq t) = \mathbb{P}(\min\{x, y \leq t\})$$

$$= 4 - \mathbb{P}(\min\{x, y \leq t\})$$

$$= 4 - \mathbb{P}(\min\{x, y \leq t\})$$

$$= 4 - \mathbb{P}(x > t, y > t)$$

$$= 4 - \mathbb{P}(x > t) \mathbb{P}(y > t)$$

$$= 4 - \mathbb{P}(x > t) \mathbb{P}(y > t)$$

$$= 4 - \mathbb{P}(x + \lambda_{z}) + \mathbb{E}(x = t)$$

$$= 4 - \mathbb{E}(x + \lambda_{z}) + \mathbb{E}(x = t)$$

(b) (5 points) Is $W := \max\{X, Y\}$ an exponential random variable? Justify your answer.

No.

$$F_{w}(t) = P(\max\{\chi, \Psi \leq t\})$$

$$= P(\chi \leq t) P(\Psi \leq t)$$

$$= (1 - e^{-\lambda_{1}t})(1 - e^{-\lambda_{2}t})$$

$$= 1 - e^{-\lambda_{1}t} \quad \text{for any } \lambda > 0.$$

5. (8 points) The life of a certain type of car battery X is normally distributed with mean 50 and the standard deviation 6 (in 1000 miles), that is, $X \in N(50, 6^2)$. Given that it has survived 50,000 miles, what is the conditional probability that the car battery survives another 15,000 miles?

$$P(X > 50 + 15 | X > 50) = \frac{P(Z > 2.5)}{P(Z > 0)}$$

$$X = 6Z + 50 = 2(1 - P(Z \le 2.5))$$

$$Z \sim N(0, 1) = 2(1 - 0.9938)$$

$$= 0.0124$$

6. Let X, Y be jointly continuous random variables with joint density

$$f_{X,Y}(x,y) = \begin{cases} \frac{1}{x^2y^2}, & x,y \geqslant 1\\ 0, & \text{otherwise.} \end{cases}$$

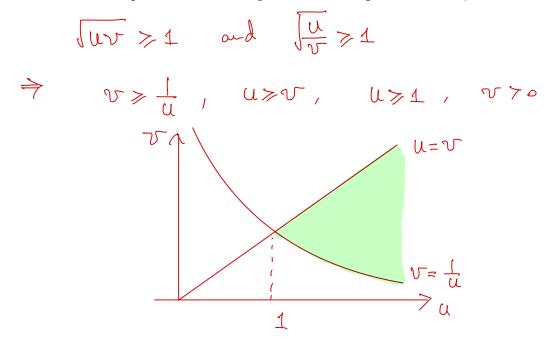
Let U, V be defined by

$$U = XY, \qquad V = X/Y.$$

(a) (10 points) Find the joint density $f_{U,V}(u, v)$ of U and V.

$$\begin{aligned} & \chi = \left\{ \overline{UV} = h_{1}(U, V) \right\} \\ & Y = \left\{ \overline{U} = h_{2}(U, V) \right\} \\ & \left[J_{h} \right] = \left| \left(\frac{\frac{1}{2} \overline{UV}}{\frac{1}{2} \overline{UV}} - \frac{1}{2} \overline{UV} \right) \right| = \frac{1}{2V} \\ & \left[\frac{1}{2} \overline{UV} - \frac{1}{2} \overline{UV} \right] \\ & \Rightarrow \quad f_{U,V}(u, v) = \quad f_{X,Y}(Uvv, \overline{Uv}, \overline{Uv}) \cdot \frac{1}{2V} \\ & = \left\{ \begin{array}{c} \frac{1}{2u^{2}v} & if \quad \overline{Uvr} > 1 \\ 0 & \text{otherwise} \end{array} \right. \end{aligned}$$

(b) (Extra credit: 4 points) Draw the region in the u - v plane where $f_{U,V}(u, v) > 0$.



(c) (Extra credit: 4 points) Find the marginal density $f_U(u)$ of U.

Method 1:

$$f_{\nabla}(u) = \int_{R} f_{\nabla, V}(u, v) dv = \int_{U} \int_{U} \frac{1}{2u^{2}v} dv , \quad u \neq 1$$

$$= \int_{U} \frac{\log u}{u^{2}} , \quad u \neq 1$$

$$= \int_{U} \frac{\log u}{u^{2}} , \quad u \neq 1$$

Method 2: For $t \ge 1$, Function = $P(XY \le t) = \int_{1}^{t} \int_{1}^{t} \frac{1}{x^{2}y^{2}} dy dx = \int_{1}^{t} \frac{1}{x^{2}} (1 - \frac{x}{t}) dx$ $= (1 - \frac{1}{t}) - \frac{1}{t} \log t$ $\int_{1}^{t} \frac{1}{x^{2}} (1 - \frac{x}{t}) dx$ $= (1 - \frac{1}{t}) - \frac{1}{t} \log t$ $\int_{1}^{t} \frac{1}{t^{2}} (1 - \frac{x}{t}) dx$ $= (1 - \frac{1}{t}) - \frac{1}{t} \log t$ $\int_{1}^{t} \frac{1}{t^{2}} (1 - \frac{x}{t}) dx$