## Practice Problems for Exam 2 (Combined)

MATH 3215, Spring 2024

## Chapter 3

- 1. The PDF of Y is given by  $f_Y(y) = \frac{c}{y^3}$  for  $1 < y < \infty$  and otherwise 0.
  - (a) Find the value of c so that  $f_Y$  is a PDF.
  - (b) Compute  $\mathbb{E}[Y]$ .

(a) 
$$\int_{1}^{\infty} f_{Y}(y) dy = \int_{1}^{\infty} \frac{c}{y^{3}} dy = \left[ -\frac{c}{2} y^{2} \right]_{1}^{\infty} = \frac{c}{2} = 1$$
  
(b) 
$$\mathbb{E}[Y] = \int_{1}^{\infty} y \cdot f_{Y}(y) dy = 2 \int_{1}^{\infty} y^{-2} dy$$
  

$$= 2 \left[ -y^{-1} \right]_{1}^{\infty} = 2.$$

2. Let X be a continuous random variable with the PDF

$$f_X(x) = \begin{cases} x+1, & -1 < x < 0, \\ 1-x, & 0 \le x < 1. \end{cases}$$

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Find the CDF of X. Draw the graph of the CDF.

$$F_{X}(x) = \int_{-\infty}^{\infty} f_{x}(t) dt$$

$$If \quad x \leq -1 \quad , \quad F_{X}(x) = 0$$

$$If \quad -1 < x < 0 \quad , \quad F_{X}(x) = \int_{-1}^{x} (t+1) dt = \frac{1}{2} (x+1)^{2}$$

$$If \quad 0 \leq x < 1 \quad , \quad F_{X}(x) = \frac{1}{2} + \int_{0}^{\infty} (1-t) dt = \frac{1}{2} + x - \frac{1}{2} x^{2}$$

$$= (-\frac{1}{2}(x-1)^{2}$$

$$If \quad x \neq 1 \quad , \quad F_{X}(x) = 1$$

- 3. Let X be the total amount of medical claims (in \$ 100,000) of the employees of a company. Assume that the PDF  $f(x) = 30x(1-x)^4$  for 0 < x < 1.
  - (a) Find  $\mathbb{E}[X]$  and  $\operatorname{Var}(X)$ .
  - (b) Find the probability that the total exceeds \$20,000.

$$\mathbb{E}[X] = \int_{0}^{1} 3p \, x^{2} \, (1-\chi)^{4} \, d\chi = \int_{0}^{1} 3p \, (21-3p+1p) = \frac{2}{7} \, d\chi$$

$$\mathbb{E}[X^{2}] = \int_{0}^{1} 3p \, x^{3} \, (1-\chi)^{4} \, d\chi = \frac{3}{28}$$

$$V_{arr}(\chi) = \mathbb{E}[\chi^{2}] - \mathbb{E}[\chi]^{2} = \frac{3}{28} - \frac{4}{49} = \frac{5}{196}$$

$$\mathbb{P}(\chi > \frac{1}{5}) = \int_{\frac{1}{5}}^{1} 3p \, \chi \, (1-\chi)^{4} \, d\chi = 3p \int_{0}^{\frac{4}{5}} (1-\xi)^{4} \, d\xi$$

$$= 3p \int_{\frac{1}{5}}^{1} 3p \, \chi \, (1-\chi)^{4} \, d\chi = 3p \int_{0}^{\frac{4}{5}} (1-\xi)^{4} \, d\xi$$

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4. The lifetime in hours of an electronic tube is a random variable having a probability density function given by

$$f(x) = xe^{-x}$$

for  $0 \le x < \infty$ . Find the expected lifetime of such a tube.

$$\mathbb{E}[X] = \int_{0}^{\infty} x^{2} e^{-X} dx = \Gamma(3) = 2! = 2.$$

- 5. Let X be the number of alpha particle emissions of carbon-14 that are counted by a Geiger counter each second. Assume that the distribution of X is Poisson with mean 5. Let W be the time in seconds before the third count is made.
  - (a) What is the distribution of W?
  - (b) Find  $\mathbb{P}(W \leq 1)$ .

(a) Since X is Poisson and W is the  
whiting time for the 2<sup>nd</sup> occurrence,  
W ~ Gamma 
$$(2, \lambda)$$
  
(b)  $P(W \leq 1) = P(X \geq 2) = 1 - P(X = 0) - P(X = 1)$   
the product when the second is th

6. A loss (in \$ 100,000) due to fire in a building has a PDF  $f(x) = \frac{1}{6}e^{-x/6}$ ,  $0 < x < \infty$ . Find the conditional probability that the loss is greater than 8 given that it is greater than 5.

Since 
$$x \sim E_{xp}(\frac{1}{6})$$
 and  $F_{x}(x) = 1 - e^{-\frac{x}{6}}$ ,  
 $P(x > 8(x > 5) = \frac{P(x > 8)}{P(x > 5)}$   
 $= \frac{1 - F_{x}(8)}{1 - F_{x}(5)} = \frac{e^{-\frac{8}{6}}}{e^{-\frac{5}{6}}} = e^{-\frac{1}{2}}$ 

7. Let  $X \sim N(650, 400)$ .

(a) Find  $\mathbb{P}(600 \le X < 660)$ .

(b) Find a constant c > 0 such that  $\mathbb{P}(|X - 650| \le c) = 0.95$ .

9. Let X be a random variable with the CDF

$$F(x) = \begin{cases} 0, & x < 0, \\ \frac{x^2}{4}, & 0 \le x < 1, \\ \frac{x+1}{4}, & 1 \le x < 2, \\ 1, & x \ge 2. \end{cases}$$
  
Find E[X], Var(X), P( $\frac{1}{4} < X < 1$ ), P(X = 1), P(X =  $\frac{1}{2}$ ), P( $\frac{1}{2} \le X < 2$ ).  

$$E[X] = \int_{0}^{1} x \cdot \frac{x}{2} \, dx + 1 \cdot P(x = \frac{1}{2}) + \int_{1}^{1} x \cdot \frac{1}{4} \, dx + 2 \cdot P(x = \lambda)$$
  

$$= \frac{1}{6} + \frac{1}{4} + \frac{3}{8} + \frac{1}{2} = \frac{1}{24} \cdot (4 + 6 + 9 + (2)) = \frac{31}{24}.$$
  

$$E[X] = \int_{0}^{1} \frac{x \cdot x}{2} \, dx + 1^{2} \cdot P(x = 1) + \int_{1}^{2} x \cdot \frac{1}{4} \, dx + 2^{2} \cdot P(x = \lambda)$$
  

$$= \frac{1}{8} + \frac{1}{4} + \frac{3}{12} + 1 = \frac{1}{8} (1 + 2 + 2 + 8) = \frac{13}{8}$$
  

$$Var(x) = \frac{13}{8} - (\frac{31}{24})^{2}$$
  

$$P(\frac{1}{4} < x < 1) = \frac{1}{4} - \frac{1}{4} \cdot (\frac{1}{4})^{2} = \frac{15}{64}, \quad P(x = 1) = \frac{1}{4}, \quad P(x = \frac{1}{2}) = 0$$
  

$$P(\frac{1}{4} \le x < \lambda) = \frac{3}{4} - \frac{1}{4} \cdot (\frac{1}{4})^{2} = \frac{11}{16}.$$

10. A loss X on a car has an exponential distribution with a mean of \$5000. If the loss X on a car is greater than the deductible of \$500, the difference X - 500 is paid to the owner of the car. Considering zero (if X < 500) as a possible payment, find the expectation of the payment.

$$P_{ayment} = u(x) = \begin{cases} x - 500 & \text{if } x \ge 500 \\ 0 & \text{otherwise} \end{cases}$$

$$E[u(x)] = \int_{too}^{\infty} (x - 500) \frac{1}{5000} e^{-\frac{x}{5000}} dx$$

$$= 5000 \cdot \int_{-\frac{1}{5}}^{\infty} (t - \frac{1}{5}) e^{-t} dt \quad (\frac{x}{5000} = t)$$

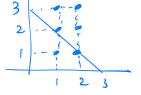
$$= 5000 \cdot \int_{0}^{\infty} u e^{-u - \frac{1}{5}} du \quad (t - \frac{1}{5} = u)$$

$$= 5000 \cdot e^{-\frac{1}{5}}.$$

## Chapter 4

1. Let X and Y be discrete random variables with joint PMF

$$f(x,y) = \frac{x+y}{21}, \qquad x = 1, 2, \quad y = 1, 2, 3.$$



- (a) Find the marginal PMFs of X and Y.
- (b) Find  $\mathbb{P}(X + Y \leq 3)$ .

$$\begin{aligned} f_{X}(x) &= \sum_{\substack{\gamma=1 \\ \gamma=1}}^{3} f(x,\gamma) = \frac{1}{21} (3x + 6) = \frac{x+2}{7} \\ f_{Y}(y) &= \sum_{\substack{x=1 \\ x=1}}^{2} f(x,\gamma) = \frac{1}{21} (3 + 2\gamma) \\ \mathbb{P}(x+Y \leq 3) &= f(1,1) + f(1,2) + f(2,1) \\ &= \frac{1}{21} ((1+1) + (1+2) + (2+1)) = \frac{8}{21} \end{aligned}$$

2. Let X and Y be discrete random variables with joint PMF

$$f(0,0) = f(1,2) = 0.2, \quad f(0,1) = f(1,1) = 0.3.$$

- (a) Find Cov(X, Y).
- (b) Find the least square regression line.

$$\begin{aligned} f_{X}(o) &= f_{X}(i) = \frac{1}{2} & f_{Y}(o) = f_{Y}(2) = o.2, \ f_{Y}(i) = o.6 \\ \mu_{X} &= o.f_{X}(o) + 1.f_{X}(i) = \frac{1}{2}, \ \mu_{Y} &= o.f_{Y}(o) + 2.f_{Y}(2) + (f_{Y}(i)) \\ \mathbb{E}[x^{1}] &= o.f_{X}(o) + 1.f_{X}(i) = \frac{1}{2}, \ \mathbb{E}[Y^{2}] = o.f_{Y}(o) + 2.f_{Y}(2) + 1.f_{Y}(i) \\ \tau_{X}^{2} &= \frac{1}{2} - (\frac{1}{2})^{2} = \frac{1}{4}, \ \sigma_{Y}^{2} &= (.4 - i^{2} = o.4) \\ \mathbb{E}[xY] &= 2.x^{1} \times y \ f(x,y) = 1.2 \cdot o.2 + (.1 \cdot o.3) = 0.7 \\ Cov(x,y) &= 0.7 - M_{X}M_{Y} = 0.2 \\ \mathbb{P} = \frac{Cov(X,Y)}{\sigma_{X}} = 2.\sqrt{\frac{5}{2}} \cdot o.2 = \sqrt{\frac{2}{5}} \\ &= (\frac{1}{5} \cdot (\frac{1}{5} \cdot 2(x - \frac{1}{2}) + 1) \\ &= (\frac{1}{5} \cdot (\frac{1}{5} \cdot 2(x - \frac{1}{2}) + 1) \\ &= (\frac{1}{5} \times (\frac{1}{5} + 2(x - \frac{1}{2}) + 1) \\ \end{bmatrix}$$

3. Let X and Y be two discrete random variables with joint PMF

$$p_{X,Y}(1,1) = p_{X,Y}(1,2) = \frac{1}{3}, \qquad p_{X,Y}(2,1) = p_{X,Y}(2,2) = \frac{1}{6}.$$

Find Cov(X, Y) and  $\rho$ . Are they independent?

$$f_{X}(1) = \frac{2}{3} \quad f_{X}(2) = \frac{1}{3} \qquad f_{Y}(1) = f_{Y}(2) = \frac{1}{2}$$

$$M_{X} = \frac{2}{3} + \frac{2}{3} = \frac{4}{3} \qquad M_{Y} = \frac{1}{2}(1+2) = \frac{3}{2}$$

$$\sigma_{X}^{2} = 2 - \frac{16}{9} = \frac{2}{9} \qquad \sigma_{Y}^{2} = \frac{5}{2} - \frac{9}{4} = \frac{1}{4}$$

$$\mathbb{E}[XY] = 1 \cdot 1 \cdot \frac{1}{3} + 1 \cdot 2 \cdot \frac{1}{3} + 2 \cdot 1 \cdot \frac{1}{6} + 2 \cdot 2 \cdot \frac{1}{6} = \frac{1}{6} \cdot (2 + 4 + 2 + 4)$$

$$= 2$$

$$C_{0}(X,Y) = 2 - \frac{4}{3} \cdot \frac{3}{2} = 0 = \beta \quad \text{Indep} \quad .$$

4. Let X be an exponential random variable with parameter 1, i.e., its PDF is given by  $f_X(x) = e^{-x}$ , x > 0. Find the PDF of  $Y = X^4$ . Are X and  $Y^{\frac{1}{2}}$  negatively correlated?

$$F_{w} \quad y \neq 0, \qquad F_{Y}(y) = P(Y \leq y) = P(X \leq y^{\frac{1}{2}})$$

$$= F_{x}(y^{\frac{1}{2}}) + \frac{1}{2} + \frac{1}{2}^{\frac{1}{2}}$$

$$= F_{x}(y^{\frac{1}{2}})$$

$$= F_{x}(y^{\frac{1}{2}}) = E[X \cdot Y^{\frac{1}{2}}] - E[X] E[Y^{\frac{1}{2}}]$$

$$= E[X^{\frac{3}{2}}] - E[X] E[X^{\frac{1}{2}}]$$

$$E[X^{\frac{1}{2}}] = \int_{0}^{\infty} x^{\frac{1}{2}} e^{-x} dx = \Gamma(\frac{1}{2}+1) = \frac{1}{2}i$$

$$G_{v}(X, Y^{\frac{1}{2}}) = 3i - 1i \cdot 2i = 4 \neq 0.$$
Positively correlated.

5. Let X and Y be discrete random variables with joint PMF  $f(x,y)=\frac{1}{9}$  for  $(x,y)\in S$  where

$$S = \{(x, y) : 0 \le x \le 2, x \le y \le x + 2, x, y \text{ are integers}\}.$$

Find  $f_X(x), f_{Y|X}(y|1), \mathbb{E}[Y|X=1]$ , and  $f_Y(y)$ .

$$f_{X}(x) = \frac{1}{3} \quad \text{for } x = 0, 1, 2$$

$$f_{Y}(x) = \frac{1}{3} \quad \text{for } y = 1, 2, 3$$

$$f_{Y}(x) = \frac{1}{3} \quad \text{for } y = 1, 2, 3$$

$$E[Y|X = 1] = 2$$

$$f_{Y}(y) = \int_{q}^{q} \quad \text{for } y = 0, 4$$

$$\int_{q}^{2} \quad \text{for } y = 1, 3$$

$$\int_{q}^{3} \quad \text{for } y = 2$$

6. Suppose that X has a geometric distribution with parameter p. and suppose the conditional distribution of Y, given X = x, is Poisson with mean x. Find  $\mathbb{E}[Y]$  and  $\operatorname{Var}(Y)$ .

$$E[Y] = E[E[Y|X]] = E[X] = \frac{1}{P}$$

$$Var(Y) = E[Var(Y|X)] + Var(E[Y|X])$$

$$= E[X] + Var(X)$$

$$= \frac{1}{P} + \frac{1-P}{P^2} = \frac{1}{P^2}$$

7. Let X and Y be two discrete random variables with joint PMF  $p_{X,Y}(x,y) = \frac{x+y}{27}$ , where x = 2,3 and y = 1, 2, 3. Find the marginal PMFs of X. Find the conditional expectation of X given Y = 2.

$$f_{X}(x) = \int_{Y=1}^{3} \frac{x+y}{27} = \frac{1}{27}(3x+6) = \frac{x+3}{9}$$

$$f_{Y}(y) = \int_{X=2}^{3} \frac{x+y}{27} = \frac{1}{27}(5+2y)$$

$$f_{X|Y}(x|2) = \left(\frac{x+2}{27}\right) \cdot 3 = \frac{x+2}{9}$$

$$\mathbb{E}[x(Y=2] = \int_{X=2}^{3} x \cdot \frac{(x+2)}{9} = \frac{1}{9}(2\cdot4+3\cdot5) = \frac{23}{9}$$

8. Let X and Y have the joint PDF

$$f(x,y) = \frac{4}{3}, \qquad 0 < x < 1, \quad x^3 < y < 1$$

and 0 otherwise. Find  $\mathbb{P}(X > Y)$ .

$$\mathbb{P}(x > Y) = \int_{0}^{1} \int_{x^{3}}^{x} \frac{4}{3} dy dx = \frac{4}{3} \int_{0}^{1} (x - x^{3}) dx$$
$$= \frac{4}{3} \cdot (\frac{1}{2} - \frac{1}{4}) = \frac{1}{3}$$

- 9. Let X be a uniform random variable over (0,2) and Y given X = x be  $U(0,x^2)$ .
  - (a) Find the joint PDF f(x, y) and marginal f<sub>Y</sub>(y).
    (b) Find E[Y|X] and E[X|Y].

Find 
$$E[Y|X]$$
 and  $E[X|Y]$ .  

$$f(X,Y) = f_{Y|X} (Y|X) \cdot f_{X}(X)$$

$$= \frac{1}{X^{2}} \cdot \frac{1}{2} \quad f_{X} \quad 0 < X < 2$$

$$f_{Y}(Y) = \int_{XY}^{2} f_{X}(Y) \, d_{X} = \frac{1}{2} \left[ -\frac{1}{X} \frac{1}{Y}^{2} = \frac{1}{2} \left( \frac{1}{XY} - \frac{1}{2} \right) \right]$$

$$E[Y|X] = \frac{1}{2} \times^{2}$$

$$f_{X|Y} (X|Y) = \frac{1}{X^{2}} \cdot \left( \frac{1}{XY} - \frac{1}{2} \right)^{1} \quad f_{X} \quad XY < 2$$

$$E[X|Y] = \int_{XY}^{2} \times \cdot \frac{1}{X^{2}} \cdot \left( \frac{1}{XY} - \frac{1}{2} \right)^{1} \, d_{X}$$

$$= \left( \frac{1}{XY} - \frac{1}{2} \right)^{1} \cdot \ln \frac{2}{XY}$$

$$E[X|Y] = \left( n \left( \frac{2}{XY} \right) \cdot \left( -\frac{2XY}{2} \right) \right].$$

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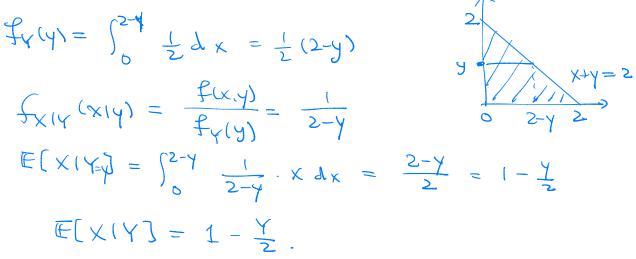
10. Let X and Y be two random variables with joint PDF  $f_{X,Y}(x,y) = 2e^{-x-y}$  for  $0 \le x \le y < \infty$ . Find the marginal PDFs of X and Y. Are they independent?

$$f_{X}(x) = \int_{X}^{\infty} 2e^{x-y} dy = 2e^{2x}, \quad x \geqslant 0$$
  

$$f_{Y}(y) = \int_{0}^{Y} 2e^{-x-y} dx = 2e^{x}(1-e^{x}), \quad y \geqslant 0$$
  

$$f_{X}(x) \cdot f_{Y}(y) \neq f(x,y) \quad \text{(Not Indep.)}$$

11. Let X and Y be two random variables with joint PDF  $f_{X,Y}(x,y) = \frac{1}{2}$  for 0 < x + y < 2, x > 0, y > 0. Find the conditional expectation  $\mathbb{E}[X|Y]$ .



12. Let (X, Y) be a bivariate normal random vector. Both X and Y have mean 1 and variance 4, while the correlation coefficient of X and Y is  $\rho = \frac{1}{4}$ . Find  $\operatorname{Var}(-2X + Y)$ . Assume now that (X, Y) is a bivariate normal vector, with X and Y having mean 9 and variance 9, but that the correlation coefficient has changed and is now given by  $\rho = 0$ . Find  $\mathbb{P}(3 \le X \le 15, Y \le 9)$  and using the tables, find an approximate value for it.

$$V_{ar}(-2x+y) = 4 V_{ar}(x) - 4 Cov(x,y) + V_{ar}(y)$$
  
= 4.4 - 4.p.  $V_{ar}(x)V_{al}(y) + 4$   
= 16 - 4. $\frac{1}{4}$ . 4 + 4 = 16.  
(x,y) Bivariate Normal and p=0 Tomphies  
X, y are Todep.  
 $\Rightarrow P(3 \le x \le 15, y \le 9)$   
=  $P(3 \le x \le 15) P(y \le 9)$   
=  $P(\frac{3-9}{3} \le z \le \frac{15-9}{3}) P(z \le \frac{9-9}{3})$   
=  $(\Xi(z) - \Xi(-2)) \cdot \Xi(0) = \frac{1}{2}(2\Xi(z)-1)$ 

- 13. The life of a certain type of automobile tire is normally distributed with mean 34,000 miles and standard deviation 4000 miles.
  - (a) What is the probability that such a tire lasts more than 40,000 miles?
  - (b) What is the probability that it lasts between 30,000 and 35,000 miles?
  - (c) Given that it has survived 30,000 miles, what is the conditional probability that the tire survives another

X = The lifetime of a tire ~ N(34,000, 4000) $(\alpha) \mathbb{P}(x > 4^{\circ}, \circ \circ \circ) = \mathbb{P}\left(\frac{x - \mu}{\sigma} > \frac{4^{\circ}, \circ \infty - \mathbf{84}, \circ \circ \circ}{4^{\circ}, \circ \circ}\right)$  $= 4 - \overline{\mathbf{E}}(1.5) = 1 - 0.933 \lambda = 0.0668$ P(30,000 (X (35,000) (6)  $= P\left(\frac{30,000 - 34,000}{4000} < \frac{X-M}{C} < \frac{3500 - 34000}{4000}\right)$  $= \overline{\mathbf{E}}(0.25) - \overline{\mathbf{E}}(-1) = \overline{\mathbf{E}}(0.25) + \overline{\mathbf{E}}(1) - 1 = 0.5987 + 0.8413 - 1$ (c)  $\mathbb{P}(X > 40,000 | X > 3000) = \frac{\mathbb{P}(X > 40000)}{\mathbb{P}(X > 30000)} = \frac{0.0668}{0.0413}$ 

b. A certain type of electrical motors is defective with probability 1/100. Pick 1000 motors and let X be the number of defective ones among these 1000 motors. Using a normal approximation, (with/without) mid-point correction, write down an expression for the probability that among the 1000 motors 13 or less are defective.

## Chapter 5

1. Let X be a uniform random variable on (-1,3) and  $Y = X^2$ . Find the PDF of Y.

Case 1:	X < 0 ,	$F^{\lambda}(\lambda) = \mathbb{B}(\lambda \langle \lambda \rangle = 0$
Care 2 :	o ≤ y < 1	$F_{Y}(y) = P(-Iy \le X \le Iy) = \frac{1}{4} \cdot 2Iy = \frac{Jy}{2}$
Case 3 >	15459	$F_{Y}(y) = \mathbb{P}(-1 \le \times \le \sqrt{y}) = \frac{1}{4}(\sqrt{y}+1)$
Case 4 :	4 7 q	$F_{Y}(y) = 1$
$f_{\alpha}(u) =$	d T (1) -	

$$f_{Y}(y) = \frac{4}{4y} F_{Y}(y) = \begin{cases} \frac{4}{5y} & 0 \le y \le 1 \\ \frac{1}{85y} & 1 \le y \le q \\ 0 & 0 \end{cases}$$

2. Let X be a random variable with PDF given by

$$f_X(x) = x^2 + \frac{10x^4}{3}$$

for 0 < x < 1, and otherwise  $f_X(x) = 0$ . Find the CDF and the PDF of  $Y = \log X$ .

Since 
$$0 < x < 1$$
,  $-\infty < Y = \log x < 0$ . For  $y < 0$ ,  
 $F_Y(y) = P(Y \in y) = P(\log x \leq y)$   
 $= P(x \leq e^y) = F_x(e^y)$   
 $f_Y(y) = f_x(e^y) \cdot e^y = (e^{2y} + \frac{10}{3}e^{4y}) \cdot e^y$  for  $y < 0$   
 $F_x(x) = \begin{cases} 0 & 7f & x \leq 0 \\ \frac{1}{3}x^3 + \frac{2}{3}x^5 & 7f & 0 < x < 1 \\ 1 & 7f & x > 1 \end{cases}$   
 $F_Y(y) = \begin{cases} \frac{1}{3}e^{3y} + \frac{2}{3}e^{5y} & 7f & y < 0 \\ \frac{1}{3}x^3 + \frac{2}{3}e^{5y} & 7f & y < 0 \end{cases}$ 

3. Let X be a random variable with density

$$f(x) = \frac{1}{\pi(1+x^2)}, \quad -\infty < x < \infty.$$

Find the density of 1/X.

$$Y = \frac{1}{x}$$

$$F_Y(y) = \mathbb{P}\left(\frac{1}{x} \leq y\right) = \mathbb{P}\left(x \geq \frac{1}{y}\right) = 1 - F_x(\frac{1}{y})$$

$$f_Y(y) = -f_x(\frac{1}{y}) \cdot \left(-\frac{1}{y^2}\right) = -\frac{1}{\pi(1+\frac{1}{y^2})} \cdot \left(-\frac{1}{y^2}\right) = \frac{1}{\pi(1+y^2)}$$

$$= f_x(y)$$