

Chapter 2. Discrete Distributions

Math 3215 Spring 2024

Georgia Institute of Technology

Section 1.
Random Variables of the Discrete
Type

Random variables

Definition

Given a random experiment with a sample space S , a function X that assigns one and only one real number $X(s) = r$ to each element in S is called a **random variable**.

The space of X is the set of real numbers $\{x : X(s) = x, s \in S\}$ and denoted by $S(X)$. (= the support of X)

Example $S = \{ \text{Male}, \text{Female} \}$

$X : S \rightarrow \mathbb{R} = \text{the set of real numbers}$

Male $\mapsto 1$

Female $\mapsto 2$

$S(X) = \{1, 2\}$

1

$Y : S \rightarrow \mathbb{R}$

Male $\mapsto -1$

Female $\mapsto 100$

$S(Y) = \{-1, 100\}$

Random variables

Example

A rat is selected at random from a cage and its sex is determined.

The set of possible outcomes is female and male. Thus, the sample space is $S = \{\text{female, male}\}$.

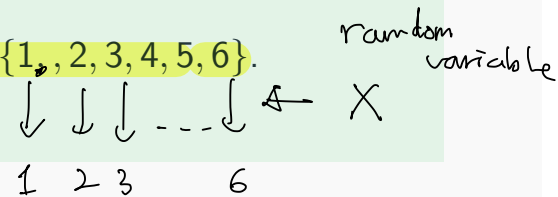
Random variables

Example

Consider a random experiment in which we roll a six-sided die.

The sample space associated with this experiment is $S = \{1, 2, 3, 4, 5, 6\}$.

Let $X(s) = s$. Compute $\mathbb{P}(2 \leq X \leq 4)$.



$$\mathbb{P}(2 \leq X \leq 4)$$

$$= \mathbb{P}(X=2 \text{ or } 3 \text{ or } 4)$$

$$= \mathbb{P}(X=2) + \mathbb{P}(X=3) + \mathbb{P}(X=4)$$

$$= \frac{1}{6} + \frac{1}{6} + \frac{1}{6} = \frac{1}{2}$$

Discrete random variables

Example uncountable set, $S = [0, 1]$

Definition

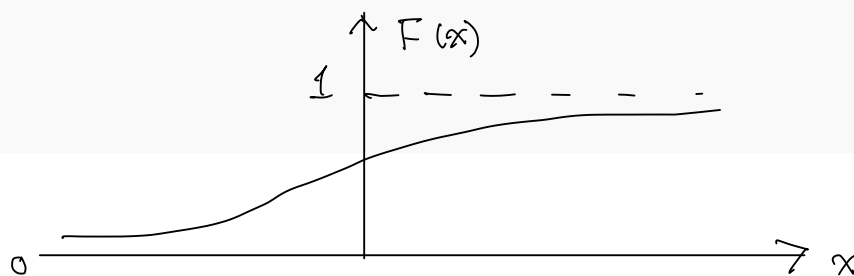
Let X be a random variable defined on a sample space S .

If S consists of $\{1, 2, 3\}$ finite outcomes or $\{1, 2, 3, \dots\}$ countable outcomes, then X is called a **discrete random variable**.

The **probability mass function (pmf)** of X is $f : \mathbb{R} \rightarrow \mathbb{R}$
 only for discrete RVs $f(x) = P(X = x)$

The **cumulative distribution function (cdf)** of X is

for any RVs, non decreasing $F : \mathbb{R} \rightarrow \mathbb{R}$



$$F(x) = P(X \leq x)$$

$$(\leq \text{ " " } \leq)$$

Discrete random variables

Properties of PMF

The pmf $f(x)$ of a discrete random variable X is a function that satisfies the following properties:

- $f(x) \geq 0$ for all x ,
- $\sum_{x \in \mathcal{S}(X)} f(x) = 1$, and
- $\mathbb{P}(X \in A) = \sum_{x \in A} f(x)$.

$$f(x) = \mathbb{P}(X=x) \geq 0$$


$$= \mathbb{P}(\Omega)$$

Discrete random variables

$$S = \{1, 2, 3, 4, 5, 6\}$$

$$\begin{array}{cccccc} \downarrow & \downarrow & \downarrow & \downarrow & \downarrow & \downarrow & X \\ 1 & 2 & 3 & 4 & 5 & 6 & \end{array}$$

Example

Roll a die, let X be the outcome.

Find the pmf and the cdf of X .

$$S(x) = \{1, 2, \dots, 6\} = S$$

PMF $f(x) = P(X=x) = \begin{cases} \frac{1}{6}, & x=1, 2, 3, 4, 5, 6 \\ 0 & \text{otherwise} \end{cases}$

$$\left(\begin{array}{l} f(1) = P(X=1) = \frac{1}{6} \\ f(0) = P(X=0) = 0 \end{array} \right)$$

CDF $F(x) = P(X \leq x) = \begin{cases} 0 & \text{for } x < 1 \\ \frac{1}{6} & 1 \leq x < 2 \\ \frac{2}{6} & 2 \leq x < 3 \\ \frac{3}{6} & 3 \leq x < 4 \\ \frac{4}{6} & 4 \leq x < 5 \\ \frac{5}{6} & 5 \leq x < 6 \\ \frac{6}{6} & x \geq 6 \end{cases}$

$$F(1) = P(X \leq 1) = \frac{1}{6}$$

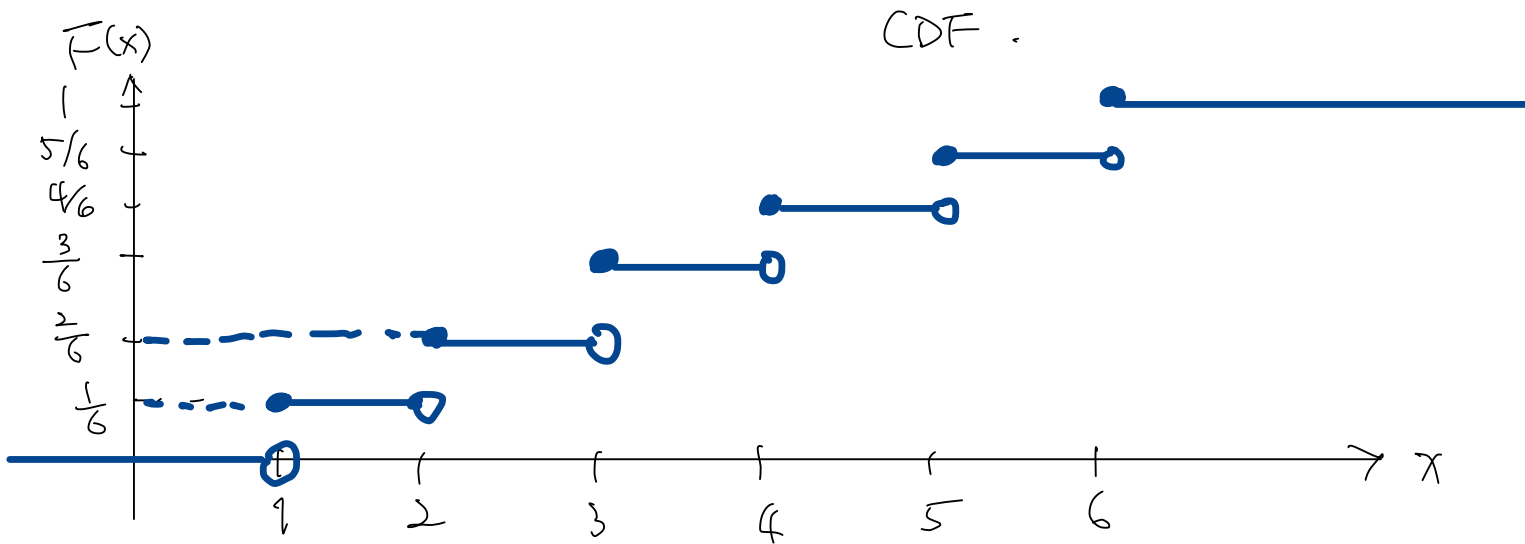
$$F(0) = P(X \leq 0) = 0$$

$$F(-1) = P(X \leq -1) = 0$$

$$F\left(\frac{3}{2}\right) = P\left(X \leq \frac{3}{2}\right) = \frac{1}{6}$$

$$F(7) = 1$$

$$F(2) = P(X \leq 2) = \frac{2}{6}$$



Discrete random variables

$\Omega = \{ (1,1), (1,2), (1,3), \dots, (2,4) \}$ 16 outcomes
 1st \rightarrow 1 2nd \rightarrow 2 3 4 ... X

Example

Roll a fair four-sided die twice.

Let X equal the larger of the two outcomes if they are different and the common value if they are the same.

Find the pmf and the cdf of X .

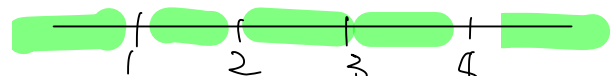
$$\mathcal{X}(x) = \{1, 2, 3, 4\}$$

PMF $f(x) = \mathbb{P}(X=x) =$

$x=1$	(1,1)	$1/16$	$x=1$
$x=2$	(1,2), (2,1)	$2/16$	$x=2$
$x=3$	(1,3), (2,3), (3,1), (3,2)	$5/16$	$x=3$
$x=4$	(2,4), (3,4), (4,1), (4,2), (4,3)	$7/16$	$x=4$
		0	otherwise

CDF $F(x) = \mathbb{P}(X \leq x)$

$x < 1$	0
$1 \leq x < 2$	$1/16$
$2 \leq x < 3$	$1/16 + 2/16$
$3 \leq x < 4$	$1/16 + 2/16 + 5/16$
$x \geq 4$	1



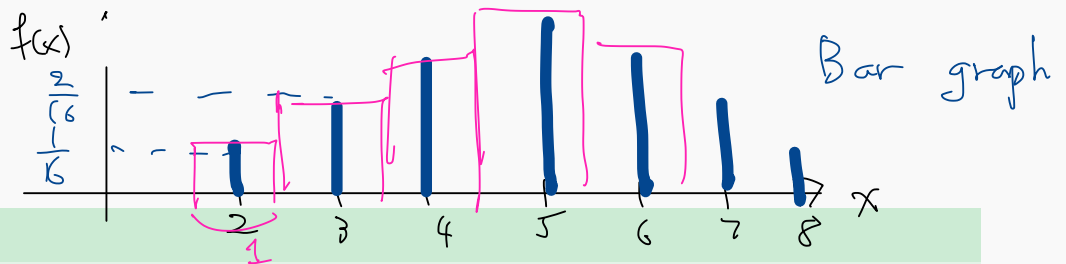
$X : \mathcal{S} \rightarrow \mathbb{R}$ RV
 \uparrow
 finite, countable Discrete RV

PMF $f(x) = f_X(x) = \mathbb{P}(X=x)$

CDF $F(x) = F_X(x) = \mathbb{P}(X \leq x)$

$\mathcal{S}(X) = \{s : X=s\}$

Bar graph, Probability histogram, relative frequency histogram



Example

A fair four-sided die with outcomes 1, 2, 3, and 4 is rolled twice.

Let X equal the sum of the two outcomes.

$\mathcal{S} = \{ (1,1), (1,2), (1,3), \dots \}$
 $\downarrow \quad \downarrow \quad \downarrow \quad \dots \quad X$
 $2 \quad 3 \quad 4$
 $\mathcal{S}(X) = \{ 2, 3, \dots, 8 \}$

PMF $f(x) = \begin{cases} 1/16 & x=2 \\ 2/16 & x=3 \quad (1,2) \quad (2,1) \\ 3/16 & x=4 \\ 4/16 & x=5 \\ 3/16 & x=6 \\ 2/16 & x=7 \\ 1/16 & x=8 \\ 0 & \text{otherwise} \end{cases}$

CDF

$$F(x) = \begin{cases} 0 & x < 2 \\ 1/16 & 2 \leq x < 3 \\ 3/16 & 3 \leq x < 4 \\ 6/16 & 4 \leq x < 5 \\ \vdots & \vdots \end{cases}$$

Bar graph, Probability histogram, relative frequency histogram

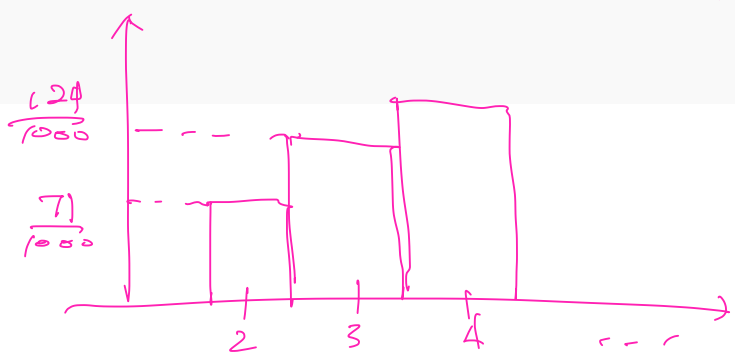
Example

Two fair four-sided dice are rolled. Write down the sum of the two outcomes. Repeat this 1000 times.

The sum of two outcomes	2	3	4	5	6	7	8
Number of Observations	71	124	194	258	177	122	54

$$\frac{71}{1000} \quad \frac{124}{1000} \quad \frac{194}{1000}$$

relative frequency



relative frequency histogram.

Section 2.

Mathematical Expectation

Definition of Expectation

Example

Consider the following game. A player roll a fair die.

If the event $A = \{1, 2, 3\}$ occurs, he receives one dollar.

If $B = \{4, 5\}$ occurs, he receives two dollars.

If $C = \{6\}$ occurs, he receives three dollars.

If the game is repeated a large number of times, what is the average payment?

Playing 6000 times

- A will happen about 3000 times
⇒ \$ 3000
- B will happen about 2000 times
⇒ \$ 4000
- C will happen about 1000 times
⇒ \$ 3000

\$ 10,000 for 6000 games

Per game, \$ $\frac{10}{6}$

$$\frac{10}{6} = \frac{1}{6} \left(\frac{3000}{6000} \cdot 1 + \frac{2000}{6000} \cdot 2 + \frac{1000}{6000} \cdot 3 \right)$$

$\frac{3000}{6000} = P(A)$ $\frac{2000}{6000} = P(B)$ $\frac{1000}{6000} = P(C)$

$$= 1 \cdot P(A) + 2 \cdot P(B) + 3 \cdot P(C)$$

$$S = \{ \underbrace{1, 2, 3}_{\downarrow}, \underbrace{4, 5}_{\downarrow}, \underbrace{6}_{\downarrow} \} \quad X$$

$$f(x) = \{ 1, 2, 3 \}$$

$$= 1 \cdot P(X=1) + 2 P(X=2) + 3 P(X=3)$$

$$= \sum_{x \in S(X)} \underbrace{x}_{=f(x)} P(X=x) = \text{Expectation of } X$$

$$= \mathbb{E}[X]$$

Definition of Expectation

Definition

If $f(x)$ is the pmf of a discrete random variable X with the space $S(X)$, and if the summation

$$\sum_{x \in S(X)} u(x)f(x)$$

exists, then the sum is called **the mathematical expectation or the expected value of $u(X)$** , and denoted by $\mathbb{E}[u(X)]$.

$$S = \{ \underbrace{1, 2, 3}_{\downarrow}, \underbrace{4, 5}_{\downarrow}, \underbrace{6}_{\downarrow} \}$$

$$\begin{array}{ccc} \cancel{1} & \cancel{2} & \cancel{3} \\ \downarrow & \downarrow & \downarrow \\ 1^2 & 2^2 & 3^2 \end{array} \quad X^2 \quad \rightarrow \quad u(x)$$

$$\mathbb{E}[X^2] = 1^2 P(X=1) + 2^2 P(X=2) + 3^2 P(X=3)$$

$$\mathbb{E}[u(x)] = u(1) P(X=1) + u(2) P(X=2) + u(3) P(X=3)$$

$$= \sum_{x \in S(X)} u(x) \cdot f(x)$$

Definition of Expectation

Example

Let the random variable X have the pmf $f(x) = \frac{1}{3}$ for $x \in \{-1, 0, 1\} = S(X)$.

Let $Y = u(X) = X^2$.

Find the pmf of Y and $\mathbb{E}[Y] = \mathbb{E}[X^2]$.

$$S(Y) = \{0, 1\}$$

$$f_Y(y) = P(Y=y) = \begin{cases} \frac{1}{3} & y=0 \\ \frac{2}{3} & y=1 \\ 0 & \text{o.w.} \end{cases}$$

$P(X=0)$ points to $\frac{1}{3}$
 $P(X=1, -1)$ points to $\frac{2}{3}$

$$\mathbb{E}[Y] = 0 \cdot f_Y(0) + 1 \cdot f_Y(1) = \frac{2}{3}$$

$$\begin{aligned} \mathbb{E}[X^2] &= (-1)^2 f_X(-1) + 0^2 f_X(0) + 1^2 f_X(1) \\ &= \frac{2}{3} \end{aligned}$$

Properties of Expectation

Theorem

1. If c is a constant, then $\mathbb{E}[c] = c$.
2. If c is a constant and u is a function, then $\mathbb{E}[cu(X)] = c\mathbb{E}[u(X)]$.
3. If c_1 and c_2 are constants and u_1 and u_2 are functions. then

$$\mathbb{E}[c_1 u_1(X) + c_2 u_2(X)] = c_1 \mathbb{E}[u_1(X)] + c_2 \mathbb{E}[u_2(X)].$$

Properties of Expectation

$$S^1(X) = \{1, 2, 3, 4\}$$

Example

Let X have the pmf $f(x) = \frac{x}{10}$ for $x = 1, 2, 3, 4$.

Find $\mathbb{E}[X]$, $\mathbb{E}[X^2]$ and $\mathbb{E}[X(5-X)]$.

$$f(x) = \begin{cases} \frac{1}{10} & x=1 \\ \frac{2}{10} & x=2 \\ \frac{3}{10} & x=3 \\ \frac{4}{10} & x=4 \end{cases}$$

$$\mathbb{E}[X] = 1 \cdot f(1) + 2 \cdot f(2) + 3 \cdot f(3) + 4 \cdot f(4)$$

$$= 1 \cdot \frac{1}{10} + 2 \cdot \frac{2}{10} + 3 \cdot \frac{3}{10} + 4 \cdot \frac{4}{10} = \frac{30}{10} = 3$$

$$\mathbb{E}[X^2] = 1^2 \cdot f(1) + 2^2 \cdot f(2) + 3^2 \cdot f(3) + 4^2 \cdot f(4)$$

$$= \frac{1^3}{10} + \frac{2^3}{10} + \frac{3^3}{10} + \frac{4^3}{10} = \frac{1}{10} \cdot (1+2+3+4)^2$$

$$= 10$$

$$\mathbb{E}[X(5-X)] = 1 \cdot (5-1) f(1) + 2 \cdot (5-2) f(2) + 3 \cdot (5-3) f(3) + 4 \cdot (5-4) f(4)$$

$$= \mathbb{E}[5X - X^2]$$

$$= \mathbb{E}[5X] - \mathbb{E}[X^2] = 5 \cdot \mathbb{E}[X] - \mathbb{E}[X^2] = 5 \cdot 3 - 10 = 5$$

Note

$$E[X^2] \neq (E[X])^2$$

$$E[u(X)] \neq u(E[X])$$

$$E[X] = \sum_{x \in \mathcal{S}(X)} x \cdot \underbrace{f(x)}_{\text{PMF, } P(X=x)}$$

$$E[u(X)] = \sum_{x \in \mathcal{S}(X)} u(x) f(x)$$

Properties of Expectation

Example

An experiment has probability of success $p \in (0, 1)$ and probability of failure $q = 1 - p$.

This experiment is repeated **independently** until the first success occurs.

Let X be the number of trials. Find $E[X]$.

$$E[X^2]$$

$$\begin{array}{ll}
 X=1 : \mathcal{S} & f(1) = P(X=1) = p \\
 X=2 : F \mathcal{S} & f(2) = P(X=2) = (1-p) \cdot p \\
 X=3 : F F \mathcal{S} & f(3) = (1-p)^2 p \\
 X=4 : F F F \mathcal{S} & f(4) = (1-p)^3 p \\
 & \vdots \\
 & \vdots
 \end{array}$$

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Geometric RV

$$A = E[X] = \underbrace{1 \cdot p}_{\text{circled}} + 2(1-p) \cdot p + 3(1-p)^2 \cdot p + 4(1-p)^3 p + \dots$$

$$-(1-p)A = \quad \quad \quad \rightarrow 1 \cdot (1-p) \cdot p \quad \rightarrow 2(1-p)^2 p \quad \rightarrow 3(1-p)^3 p + \dots$$

$$\frac{A - (1-p)A}{p \cdot A} = \underbrace{1 \cdot p}_{\text{circled}} + 1 \cdot \cancel{(1-p) \cdot p} + 1 \cdot \cancel{(1-p)^2 \cdot p} + 1 \cdot \cancel{(1-p)^3 \cdot p} + \dots$$

$$-(1-p) \cdot pA = \quad \quad \quad \cancel{(1-p) \cdot p} + \cancel{(1-p)^2 \cdot p} + \cancel{(1-p)^3 \cdot p} + \dots$$

$$pA = pA - (1-p) \cdot pA = p \Rightarrow pA = 1 \quad \therefore A = \frac{1}{p}$$

1st trial



S
F

Given this

Exp. # of trial = 1

$$F[x] = A$$

Given this

Exp # of trial
= 1 + A

$$A = p \cdot 1 + (1-p)(1+A)$$

$$A = \underbrace{p + (1-p)} + \overbrace{(1-p) \cdot A} = 1 + \cancel{A} - p \cdot A$$

$pA = 1 \quad \therefore \quad A = \frac{1}{p}$

Section 3.

Special Mathematical Expectations

Moments

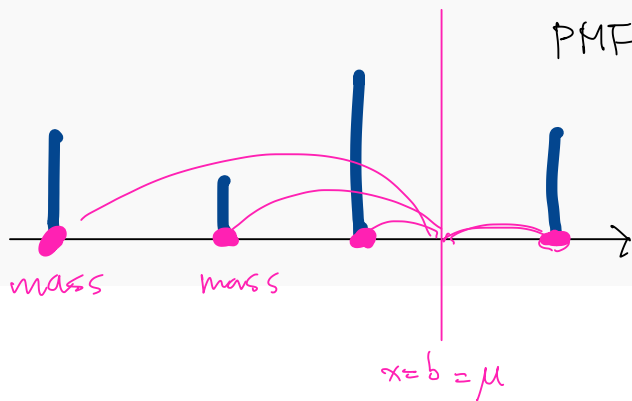
The 1st moment of X about b
 $= \mathbb{E}[(X - b)^1]$

The expectation or mean of a random variable X is

$$\mu = \mathbb{E}[X] = \sum xf(x).$$

This is also called the first moment about the origin.

The first moment about the mean μ is $\mathbb{E}[X - \mu] = \mathbb{E}[x] - \mathbb{E}[\mu] = \mathbb{E}[x] - \mu = 0$



Moments

The second moment of X about b is $\mathbb{E}[(X - b)^2]$.

If $b = \mu$, it is also called **the variance** of X and denoted by $\text{Var}(X) = \sigma^2 = \sigma_x^2$

Its positive square root is **the standard deviation** of X and denoted by $\text{Std}(X) = \sigma = \sigma_x$

$$\begin{aligned}\text{Var}(X) &= \sigma_x^2 = \mathbb{E}[(X - \mu)^2] \\ &= \mathbb{E}[(X - \mathbb{E}(X))^2]\end{aligned}$$

$$= \sqrt{\text{Var}(X)}$$

$$\begin{aligned}\text{Std}(X) &= \sigma_x = \sqrt{\text{Var}(X)} \\ &= \sqrt{\mathbb{E}[(X - \mu)^2]}\end{aligned}$$

Moments

Example

Roll a fair die and let X be the outcome.

Find $\mathbb{E}[X]$ and $\text{Var}(X)$.

$$\mathbb{E}[X] = 1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} + \dots + 6 \cdot \frac{1}{6} = \frac{1}{6} \cdot \underbrace{(1+2+\dots+6)}_{=21} = \frac{7}{2}$$

$$\text{Var}(X) = \mathbb{E}\left[\underbrace{\left(X - \frac{7}{2}\right)^2}_{=21}\right]$$

$$= \left(1 - \frac{7}{2}\right)^2 \cdot \frac{1}{6} + \left(2 - \frac{7}{2}\right)^2 \cdot \frac{1}{6} + \left(3 - \frac{7}{2}\right)^2 \cdot \frac{1}{6}$$

$$+ \left(4 - \frac{7}{2}\right)^2 \cdot \frac{1}{6} + \left(5 - \frac{7}{2}\right)^2 \cdot \frac{1}{6} + \left(6 - \frac{7}{2}\right)^2 \cdot \frac{1}{6}$$

$$= \frac{1}{6} \cdot 2 \cdot \frac{1}{2} \cdot \underbrace{(1^2 + 3^2 + 5^2)}_{=35} = \frac{35}{12}$$

$$= \mathbb{E}\left[X^2 - 2 \cdot X \cdot \frac{7}{2} + \left(\frac{7}{2}\right)^2\right]$$

$$= \mathbb{E}\left[X^2 - 7X + \frac{49}{4}\right] = \underbrace{\mathbb{E}[X^2]}_{=35} - 7 \cdot \underbrace{\mathbb{E}[X]}_{=7/2} + \frac{49}{4}$$

$$\mathbb{E}[X^2] = \frac{1}{6} \cdot 1^2 + \dots + \frac{1}{6} \cdot 6^2 = \frac{1}{6} \cdot \underbrace{(1^2 + \dots + 6^2)}$$

Moments

In general, the r -th moment of X about b is $\mathbb{E}[(X - b)^r]$.

Definition

Index of skewness is defined by

$$\gamma = \mathbb{E}[(X - \mu)^3] / \sigma^3 = \frac{\mathbb{E}[(X - \mu)^3]}{(\mathbb{E}[(X - \mu)^2])^{3/2}}$$

Moments

$$\mu = E[X] \quad \sigma^2 = E[(X - \mu)^2]$$

$$\gamma = \frac{E[(X - \mu)^3]}{\sigma^3}$$

Example

Let $f(x) = \frac{4-x}{6}$ for $x = 1, 2, 3$ be the pmf of X . Compute the index of skewness.

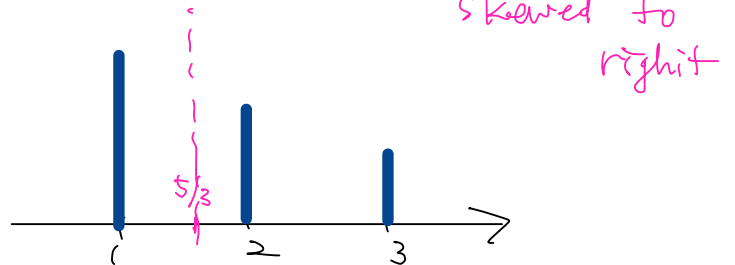
$$\mu = E[X] = 1 \cdot \frac{3}{6} + 2 \cdot \frac{2}{6} + 3 \cdot \frac{1}{6} = \frac{1}{6} \cdot (1 \cdot 3 + 2 \cdot 2 + 3 \cdot 1) = \frac{10}{6}$$

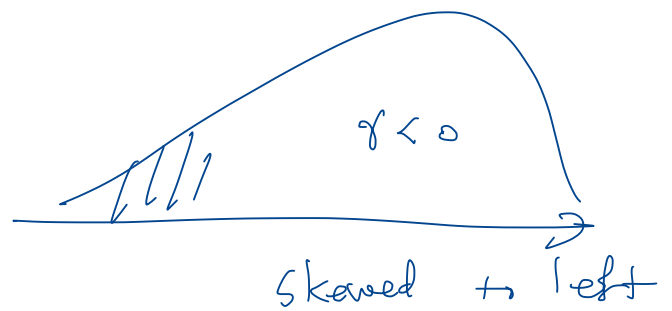
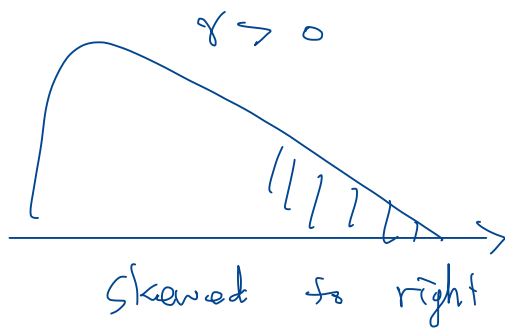
$$\begin{aligned} \sigma^2 &= E\left[\left(X - \frac{5}{3}\right)^2\right] = \underbrace{\left(1 - \frac{5}{3}\right)^2}_{\frac{4}{9}} \cdot \frac{3}{6} + \underbrace{\left(2 - \frac{5}{3}\right)^2}_{\frac{1}{9}} \cdot \frac{2}{6} + \underbrace{\left(3 - \frac{5}{3}\right)^2}_{\frac{16}{9}} \cdot \frac{1}{6} \\ &= \frac{1}{9 \cdot 6} \cdot \left((-2)^2 \cdot 3 + 1^2 \cdot 2 + 4^2 \cdot 1 \right) = \frac{20}{9 \cdot 6} = \frac{5}{9} \end{aligned}$$

$$\begin{aligned} E\left[\left(X - \frac{5}{3}\right)^3\right] &= \underbrace{\left(1 - \frac{5}{3}\right)^3}_{-\frac{8}{27}} \cdot \frac{3}{6} + \underbrace{\left(2 - \frac{5}{3}\right)^3}_{\frac{1}{27}} \cdot \frac{2}{6} + \underbrace{\left(3 - \frac{5}{3}\right)^3}_{\frac{64}{27}} \cdot \frac{1}{6} \\ &= \frac{1}{27 \cdot 6} \left((-2)^3 \cdot 3 + 1^3 \cdot 2 + 4^3 \cdot 1 \right) = \frac{7}{27} \end{aligned}$$

$-24 \quad + 2 \quad + 64 = 42$

$$\gamma = \frac{7/27}{(5/9)^{3/2}} = \frac{7}{27} \cdot \frac{27}{5 \cdot \sqrt{5}} = \frac{7}{5\sqrt{5}} > 0$$





Moments

Theorem

$$\sigma^2 = \mathbb{E}[(X - \mu)^2] = \mathbb{E}[X^2] - \mu^2 = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$$

$$= \mathbb{E}[X^2 - 2\mu X + \mu^2]$$

$$= \mathbb{E}[X^2] - 2\mu \underbrace{\mathbb{E}[X]}_{\mu} + \mu^2$$

$$= \mathbb{E}[X^2] - 2\mu^2 + \mu^2$$

$$= \mathbb{E}[X^2] - \mu^2$$

$$\text{Var}(X) = \sum_i \underbrace{(x_i - \mu)^2}_{\geq 0} \underbrace{f(x_i)}_{\geq 0} \geq 0$$

$$\mathbb{E}[X^2] - (\mathbb{E}[X])^2 \geq 0$$

$$\mathbb{E}[X^2] \geq (\mathbb{E}[X])^2$$

Jensen's Inequality

$$u(x) = e^{tx}$$

$$\mathbb{E}[u(x)] = \mathbb{E}[e^{tx}]$$

Moment generating functions

Definition

Let X be a discrete random variable and assume that there exists $h > 0$ such that

$$\mathbb{E}[e^{tX}] = \sum e^{tx} f(x) \quad \leftarrow \text{a function of } t$$

is finite for all $t \in (-h, h)$. Then, $M(t) = \mathbb{E}[e^{tX}]$ is called **the moment generating function (mgf)**.

$$M(0) = \mathbb{E}[e^{0 \cdot X}] = 1$$

$$M'(t) \Big|_{t=0} = \frac{d}{dt} \mathbb{E}[e^{tX}] \Big|_{t=0} = \mathbb{E} \left[\frac{d}{dt} e^{tX} \right] \Big|_{t=0}$$

$= X e^{tX}$ by

$$= \mathbb{E}[X e^{tX}] \Big|_{t=0}$$

$$= \mathbb{E}[X]$$

$$M''(0) = \mathbb{E}[X^2 e^{tX}] \Big|_{t=0}$$

$$= \mathbb{E}[X^2]$$

$(e^t)' = e^t$
 Chain Rule
 $(e^{f(t)})' = e^{f(t)} \cdot f'(t)$
 $= e^{f(t)} \cdot f'(t)$

Moment generating functions

Properties

1. $M(0) = 1$
2. $M'(0) = \mathbb{E}[X]$
3. $M''(0) = \mathbb{E}[X^2]$
4. In general, $M^{(r)}(0) = \mathbb{E}[X^r]$.

$$M(t) = \mathbb{E}[e^{tX}]$$

Hint: $a + a \cdot r + a \cdot r^2 + \dots = \frac{a}{1-r}$.

Moment generating functions

$X = \#$ of trials until first success
 prob = p .

Example PMF $X = 1, 2, 3, \dots$

Let $f(x) = q^{x-1}p$ where $p \in (0, 1)$ and $q = 1 - p$.

Compute $M(t)$.

$$M(t) = \mathbb{E}[e^{tX}] = e^{t \cdot 1} \cdot q^0 \cdot p + e^{2t} q^1 \cdot p + e^{3t} q^2 \cdot p + \dots$$

$\xrightarrow{\text{Factor } e^t \cdot q}$

$$= \frac{e^t \cdot p}{1 - e^t \cdot q}$$

Exercise

• $M'(0) = \frac{1}{p} = \mathbb{E}[X]$.

• $\mathbb{E}\left[\left(\frac{1}{2}\right)^X\right] = ?$

• $M(t) < \infty$? for all t ?

$$\text{Var}(X) = \text{Std}(X)^2 = \mathbb{E}[(X - \mu)^2], \quad \mu = \mathbb{E}[X]$$

$$= \mathbb{E}[X^2] - (\mathbb{E}[X])^2$$

$$M(t) = \mathbb{E}[e^{tX}] \quad -h < t < h$$

$$M(0) = 1$$

$$M'(0) = \mathbb{E}[X]$$

$$M''(0) = \mathbb{E}[X^2]$$

$$X, \quad \text{PMF} \quad f(x) = (1-p)^{x-1} \cdot p, \quad p \in (0,1)$$

$$x = 1, 2, \dots$$

$$M(t) = \mathbb{E}[e^{tX}] = \sum_{x=1}^{\infty} e^{tx} \cdot f(x)$$

$$= e^t \cdot (1-p)^0 \cdot p + e^{2t} (1-p)^1 \cdot p + e^{3t} (1-p)^2 \cdot p + \dots$$

$\underbrace{\hspace{10em}}_{x \quad e^t(1-p)} \quad \underbrace{\hspace{10em}}_{e^t(1-p)} \quad \dots$

(Geometric Series)

$$= \frac{\text{First}}{1 - \text{Ratio}} = \frac{e^t \cdot p}{1 - e^t(1-p)} \quad \triangleleft \quad t < \ln\left(\frac{1}{1-p}\right)$$

↑
only works when $|e^t(1-p)| < 1$
" $e^t(1-p)$

$$e^t < \frac{1}{1-p}$$

$$t < \ln\left(\frac{1}{1-p}\right)$$

$$\mathbb{E}[2^X] = 2^1 \cdot (1-p)^0 \cdot p + 2^2 (1-p)^1 \cdot p + 2^3 (1-p)^2 \cdot p + \dots$$

$$= \frac{2p}{1 - 2(1-p)}$$

$$E[2^X] = M(t) = E[e^{tX}] = M(\log 2)$$

$$2^X = e^{tX} \rightarrow X \cdot \log 2 = t \cdot X$$

$$t = \log 2$$

Section 4.

The Binomial Distribution

Bernoulli random variables

A **Bernoulli experiment**, more commonly called a **Bernoulli trial**, is a random experiment with two outcomes.

Say $S = \{\text{success, failure}\}$ and $\mathbb{P}(\text{success}) = p$ for some $p \in (0, 1)$. Then $\mathbb{P}(\text{failure}) = q = 1 - p$.

A random variable X is a **Bernoulli random variable** with **success probability** p is $X = 1$ if success and 0 otherwise.

$$X = \begin{cases} 1 & \text{with prob. } p \\ 0 & \text{with prob. } (1-p) \end{cases}$$

• PMF $f(x) = \begin{cases} p & , x = 1 \\ 1-p & , x = 0 \\ 0 & , \text{otherwise} \end{cases}$

25

• $\mathbb{E}[X] = 0 \cdot (1-p) + 1 \cdot p = p$

• $\text{Var}(X) = \mathbb{E}[X^2] - (\mathbb{E}[X])^2 = p - p^2 = p(1-p) = p \cdot q$

$$\mathbb{E}[X^2] = 0^2(1-p) + 1^2 \cdot p = p$$

• MGF $M(t) = \mathbb{E}[e^{tx}] = e^{t \cdot 0} \cdot (1-p) + e^{t \cdot 1} p = pe^t + (1-p)$

Bernoulli random variables

Theorem

Let X be a Bernoulli random variable with success probability p .

$$\mathbb{E}[X] = p$$

$$\text{Var}[X] = p(1-p)$$

Binomial random variables

Consider a sequence of independent Bernoulli experiments with success probability p .

Let X be the number of success trials in the first n experiments.

This is called a **Binomial random variable** with the number of trials n and success probability p .

We use the notation $X \sim b(n, p) = \text{Bin}(n, p)$.

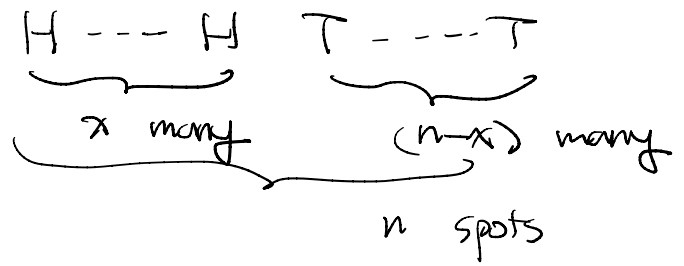
Ex $n = 5$ $p = \frac{1}{2}$ $X \sim \text{Bin}(5, \frac{1}{2})$
 $X = \#$ of Heads in 5 exp.

• PMF .

$f(x) =$	$\left(\frac{1}{2}\right)^5$,	$x=0$	TTTTT	27
$\binom{5}{1} = 5C_1$	$5 \left(\frac{1}{2}\right)^5$,	$x=1$	HTTTTT, THTTTT,	
$5C_2 =$	$\binom{5}{2} \left(\frac{1}{2}\right)^5$	$x=2$		
	$\binom{5}{3} \left(\frac{1}{2}\right)^5$	$x=3$		
	$\binom{5}{4} \left(\frac{1}{2}\right)^5$	$x=4$		
	$\binom{5}{5} \left(\frac{1}{2}\right)^5$	$x=5$		

for $x = 0, 1, 2, \dots, n$

In general, $X \sim \text{Bin}(n, p)$, $f(x) = \binom{n}{x} p^x (1-p)^{n-x}$



Binomial random variables

Theorem

Let X a binomial random variable with the number of trials n and success probability p .

$$X \sim \text{Bin}(n, p)$$

The pmf of X is

$$f(x) = \binom{n}{x} p^x (1-p)^{n-x}, \quad x=0, 1, \dots, n$$

$$\mathbb{E}[X] = n \cdot p$$

$$\text{Var}[X] = n \cdot p \cdot (1-p)$$

$$\sum_{x=0}^n f(x) = 1 = \binom{n}{0} p^0 (1-p)^n + \binom{n}{1} p^1 (1-p)^{n-1} + \dots + \binom{n}{n} p^n (1-p)^0$$

Binomial Thm $(a+b)^n = \binom{n}{0} a^n b^0 + \binom{n}{1} a^{n-1} b^1 + \dots + \binom{n}{n-1} a^1 b^{n-1} + \binom{n}{n} a^0 b^n$

$$1 = ((1-p) + p)^n = \binom{n}{0} p^0 (1-p)^n + \binom{n}{1} p^1 (1-p)^{n-1} + \dots + \binom{n}{n} p^n (1-p)^0$$

$$\mathbb{E}[X] = \sum_{x=0}^n x \cdot f(x) = \sum_{x=1}^n x \cdot \frac{n!}{x!(n-x)!} p^x (1-p)^{n-x}$$

$$\begin{aligned} x \binom{n}{x} &= x \cdot \frac{n!}{x!(n-x)!} = \frac{n \cdot (n-1)!}{(x-1)! (n-1-(x-1))!} = n \cdot \binom{n-1}{x-1} \\ &= n \cdot \sum_{x=1}^n \binom{n-1}{x-1} p^{x-1} (1-p)^{n-x} = (1-p)^{n-1} \sum_{x=1}^n \binom{n-1}{x-1} p^{x-1} (1-p)^{n-x} \end{aligned}$$

$$= n \cdot p \sum_{x=1}^n \binom{n-1}{x-1} p^{x-1} (1-p)^{(n-1)-(x-1)}$$

↓ Binomial Thm

$$= np (p + (1-p))^{n-1} = \underline{np}$$

$$\text{Var}(X) = E[X^2] - (E[X])^2 = E[X^2] - (np)^2$$

$$E[X^2] = ?$$

$E[X(X-1)]$ is "double".

$$= E[X^2 - X] = E[X^2] - \underbrace{E[X]}_{np}$$

Binomial RV Repeat Bernoulli Trials n times
indep. (2 outcomes)
 same success prob. p

$X = \#$ of success $\sim \text{Bin}(n, p)$ ($0 < p < 1$)

$$\text{PMF} = f(x) = \binom{n}{x} p^x (1-p)^{n-x} \quad (1-p = q)$$

$$E[X] = np \quad \text{Var}(X) = np(1-p) = npq$$

$$\text{MGF} = M(t) = E[e^{tx}] = (e^t p + (1-p))^n$$

MGF ① (Heuristic) $e^t = 1 + t + \frac{1}{2!} t^2 + \frac{1}{3!} t^3 + \dots$

$$E[e^{tx}] = E\left[1 + (tX) + \frac{(tX)^2}{2!} + \frac{(tX)^3}{3!} + \dots\right]$$

$$= 1 + t \cdot E[X] + \frac{t^2}{2!} E[X^2] + \frac{t^3}{3!} E[X^3] + \dots$$

$$\Rightarrow M'(0) = E[X], \quad M''(0) = E[X^2], \dots$$

② Sometimes, we define a RV $X \rightarrow$ give MGF easily
 not PMF.

$$M_X(t) \approx M_Y(t) \Rightarrow X \stackrel{d}{\approx} Y$$

③ $X+Y \xrightarrow{\text{indep.}} M_{X+Y} = M_X \cdot M_Y$

Repeat B.T. 20. Success prob. = 0.3



$X_1 = \#$ of success

$X_2 = \#$ of success

$$X_1 \sim \text{Bin}(10, 0.3)$$

$$X_2 \sim \text{Bin}(10, 0.3)$$

$$X_1 + X_2 \sim \text{Bin}(20, 0.3)$$

$$M_{X_1} = (e^t p + (1-p))^{10} = M_{X_2}$$

$$M_{X_1+X_2} = (e^t p + (1-p))^{20} = M_{X_1} \cdot M_{X_2}$$

Binomial random variables

$$X \sim \text{Bin}\left(8, \frac{1}{5}\right)$$

Example

Out of millions of instant lottery tickets, suppose that 20% are winners. If eight such tickets are purchased, what is the probability of purchasing two winning tickets?

$$P(X = 2) = f(2) = \binom{8}{2} \cdot \left(\frac{1}{5}\right)^2 \cdot \left(\frac{4}{5}\right)^6 //$$

Binomial random variables

Example

H5N1 is a type of influenza virus that causes a severe respiratory disease in birds called avian influenza (or “bird flu”).

Although human cases are rare, they are deadly; according to the World Health Organization the mortality rate among humans is 60%.

Let X equal the number of people, among the next 25 reported cases, who survive the disease.

+ same prob

Assuming independence, the distribution of X is $b(25, 0.4)$. What is the probability that ten or fewer of the cases survive?

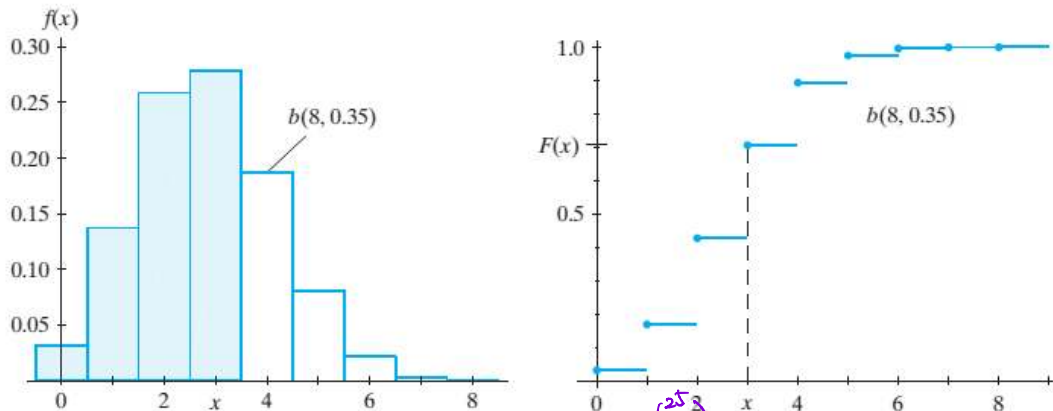
$$\begin{aligned} P(X \leq 10) &= \sum_{x=0}^{10} P(X=x) \\ &= \sum_{x=0}^{10} \binom{25}{x} (0.4)^x (0.6)^{25-x} \end{aligned}$$

$$X \sim \text{Bin}(25, 0.4)$$

$$P(X \leq 10)$$

$$n = 25, p = 0.4, x = 10$$

Table II The Binomial Distribution



$$CDF \rightarrow F(x) = P(X \leq 10) = \sum_{k=0}^{10} \frac{n!}{k!(n-k)!} p^k (1-p)^{n-k}$$

$$n=25, p=0.4, 1-p=0.6$$

n	x	p									
		0.05	0.10	0.15	0.20	0.25	0.30	0.35	0.40	0.45	0.50
2	0	0.9025	0.8100	0.7225	0.6400	0.5625	0.4900	0.4225	0.3600	0.3025	0.2500
	1	0.9975	0.9900	0.9775	0.9600	0.9375	0.9100	0.8775	0.8400	0.7975	0.7500
	2	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
3	0	0.8574	0.7290	0.6141	0.5120	0.4219	0.3430	0.2746	0.2160	0.1664	0.1250
	1	0.9928	0.9720	0.9392	0.8960	0.8438	0.7840	0.7182	0.6480	0.5748	0.5000
	2	0.9999	0.9990	0.9966	0.9920	0.9844	0.9730	0.9571	0.9360	0.9089	0.8750
	3	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
4	0	0.8145	0.6561	0.5220	0.4096	0.3164	0.2401	0.1785	0.1296	0.0915	0.0625
	1	0.9860	0.9477	0.8905	0.8192	0.7383	0.6517	0.5630	0.4752	0.3910	0.3125
	2	0.9995	0.9963	0.9880	0.9728	0.9492	0.9163	0.8735	0.8208	0.7585	0.6875
	3	1.0000	0.9999	0.9995	0.9984	0.9961	0.9919	0.9850	0.9744	0.9590	0.9375
	4	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
5	0	0.7738	0.5905	0.4437	0.3277	0.2373	0.1681	0.1160	0.0778	0.0503	0.0312
	1	0.9774	0.9185	0.8352	0.7373	0.6328	0.5282	0.4284	0.3370	0.2562	0.1875
	2	0.9988	0.9914	0.9734	0.9421	0.8965	0.8369	0.7648	0.6826	0.5931	0.5000
	3	1.0000	0.9995	0.9978	0.9933	0.9844	0.9692	0.9460	0.9130	0.8688	0.8125
	4	1.0000	1.0000	0.9999	0.9997	0.9990	0.9976	0.9947	0.9898	0.9815	0.9688
	5	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
6	0	0.7351	0.5314	0.3771	0.2621	0.1780	0.1176	0.0754	0.0467	0.0277	0.0156
	1	0.9672	0.8857	0.7765	0.6553	0.5339	0.4202	0.3191	0.2333	0.1636	0.1094
	2	0.9978	0.9842	0.9527	0.9011	0.8306	0.7443	0.6471	0.5443	0.4415	0.3438
	3	0.9999	0.9987	0.9941	0.9830	0.9624	0.9295	0.8826	0.8208	0.7447	0.6562
	4	1.0000	0.9999	0.9996	0.9984	0.9954	0.9891	0.9777	0.9590	0.9308	0.8906
	5	1.0000	1.0000	1.0000	0.9999	0.9998	0.9993	0.9982	0.9959	0.9917	0.9844
	6	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
7	0	0.6983	0.4783	0.3206	0.2097	0.1335	0.0824	0.0490	0.0280	0.0152	0.0078
	1	0.9556	0.8503	0.7166	0.5767	0.4449	0.3294	0.2338	0.1586	0.1024	0.0625

$$P(X = 10) = P(X \leq 10) - P(X \leq 9)$$

$$(a + b)^n = \sum_{x=0}^n \binom{n}{x} a^x b^{n-x}$$

Binomial random variables

$$X \sim \text{Bin}(n, p)$$

Theorem

The mgf of a binomial random variable X is

$$M(t) =$$

$$(e^t)^x \cdot p^x = (e^t \cdot p)^x$$

$$\begin{aligned} M(t) &= \mathbb{E}[e^{tX}] = \sum_{x=0}^n \underbrace{(e^t)^x}_{\substack{\uparrow \\ \text{from } (e^t \cdot p)^x}} \binom{n}{x} \underbrace{p^x}_{\substack{\uparrow \\ \text{from } (e^t \cdot p)^x}} (1-p)^{n-x} \\ &= \sum_{x=0}^n \binom{n}{x} \underbrace{(e^t \cdot p)^x}_a \underbrace{(1-p)^{n-x}}_b \\ &= (e^t \cdot p + (1-p))^n \end{aligned}$$

Binomial random variables

Exercise

It is believed that approximately 75% of American youth now have insurance due to the health care law.

Suppose this is true, and let X equal the number of American youth in a random sample of $n = 15$ with private health insurance.

How is X distributed? Find the probability that X is at least 10. Find the mean, variance, and standard deviation of X .

$$X \sim \text{Bin}(15, 0.75)$$

① indep.

② the same distribution (identical)

$$P(X \geq 10) = \sum_{x=10}^{15} \binom{15}{x} (0.75)^x (0.25)^{15-x}$$

Use the table!

$$P(X \geq 10) = P(Y \leq 5)$$

" # of people not having insurance.
 $Y \sim \text{Bin}(15, 0.25)$

$$X \sim \text{Bin}(15, 0.75) \quad , \quad P(X \geq 10) = 1 - P(X \leq 9)$$

Table II continued		p									
n	x	0.05	0.10	0.15	0.20	0.25	0.30	0.35	0.40	0.45	0.50
15	11	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9994	0.9978	0.9935
	12	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9997	0.9991
	13	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999
	14	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
	0	0.4633	0.2059	0.0874	0.0352	0.0134	0.0047	0.0016	0.0005	0.0001	0.0000
	1	0.8290	0.5490	0.3186	0.1671	0.0802	0.0353	0.0142	0.0052	0.0017	0.0005
	2	0.9638	0.8159	0.6042	0.3980	0.2361	0.1268	0.0617	0.0271	0.0107	0.0037
	3	0.9945	0.9444	0.8227	0.6482	0.4613	0.2969	0.1727	0.0905	0.0424	0.0176
	4	0.9994	0.9873	0.9383	0.8358	0.6865	0.5155	0.3519	0.2173	0.1204	0.0592
	5	0.9999	0.9978	0.9832	0.9389	0.8516	0.7216	0.5643	0.4032	0.2608	0.1509
	6	1.0000	0.9997	0.9964	0.9819	0.9434	0.8689	0.7548	0.6098	0.4522	0.3036
	7	1.0000	1.0000	0.9994	0.9958	0.9827	0.9500	0.8868	0.7869	0.6535	0.5000
	8	1.0000	1.0000	0.9999	0.9992	0.9958	0.9848	0.9578	0.9050	0.8182	0.6964
	9	1.0000	1.0000	1.0000	0.9999	0.9992	0.9963	0.9876	0.9662	0.9231	0.8491
	10	1.0000	1.0000	1.0000	1.0000	0.9999	0.9993	0.9972	0.9907	0.9745	0.9408
11	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9995	0.9981	0.9937	0.9824	
12	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9987	0.9989	0.9963	
13	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9995	
14	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	
15	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	
16	0	0.4401	0.1853	0.0743	0.0281	0.0100	0.0033	0.0010	0.0003	0.0001	0.0000
	1	0.8108	0.5147	0.2839	0.1407	0.0635	0.0261	0.0098	0.0033	0.0010	0.0003
	2	0.9571	0.7892	0.5614	0.3518	0.1971	0.0994	0.0451	0.0183	0.0066	0.0021
	3	0.9930	0.9316	0.7899	0.5981	0.4050	0.2459	0.1339	0.0651	0.0281	0.0106
	4	0.9991	0.9830	0.9209	0.7982	0.6302	0.4499	0.2892	0.1666	0.0853	0.0384
	5	0.9999	0.9967	0.9765	0.9183	0.8103	0.6598	0.4900	0.3288	0.1976	0.1051
	6	1.0000	0.9995	0.9944	0.9733	0.9204	0.8247	0.6881	0.5272	0.3660	0.2272
	7	1.0000	0.9999	0.9989	0.9930	0.9729	0.9256	0.8406	0.7161	0.5629	0.4018
	8	1.0000	1.0000	0.9998	0.9985	0.9925	0.9743	0.9329	0.8577	0.7441	0.5982
	9	1.0000	1.0000	1.0000	0.9998	0.9984	0.9929	0.9771	0.9417	0.8759	0.7728
	10	1.0000	1.0000	1.0000	1.0000	0.9997	0.9984	0.9938	0.9809	0.9514	0.8949
	11	1.0000	1.0000	1.0000	1.0000	1.0000	0.9997	0.9987	0.9951	0.9851	0.9616
	12	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9998	0.9991	0.9965	0.9894
	13	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9994	0.9979
	14	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9997
	15	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
16	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	
20	0	0.3585	0.1216	0.0388	0.0115	0.0032	0.0008	0.0002	0.0000	0.0000	0.0000
	1	0.7358	0.3917	0.1756	0.0692	0.0243	0.0076	0.0021	0.0005	0.0001	0.0000
	2	0.9245	0.6769	0.4049	0.2061	0.0913	0.0355	0.0121	0.0036	0.0009	0.0002
	3	0.9841	0.8670	0.6477	0.4114	0.2252	0.1071	0.0444	0.0160	0.0049	0.0013
	4	0.9974	0.9568	0.8298	0.6296	0.4148	0.2375	0.1182	0.0510	0.0189	0.0059

Section 5.

The Hypergeometric Distribution

Ex 6 Red balls 4 Blue balls

Choose 4 balls at random with replacement

$X = \#$ of red balls chosen. $\sim \text{Bin}(4, 0.6)$

Question: without replacement.

$$\left(\begin{aligned} P(2^{\text{nd}} = R) &= P(BR) + P(RR) \\ &= \frac{4}{10} \cdot \frac{6}{9} + \frac{6}{10} \cdot \frac{5}{9} = \frac{24+30}{10 \cdot 9} = \frac{6}{10} \end{aligned} \right)$$

$$P(X=2) = \frac{\binom{6}{2} \cdot \binom{4}{2}}{\binom{10}{4}}$$

$$= \frac{6}{10} \cdot \frac{5}{9} \cdot \frac{4}{8} \cdot \frac{3}{7} \cdot \binom{4}{2}$$

(RRBB)

of arrangement.

The Hypergeometric Distribution

There is a collection of N_1 red balls and N_2 blue balls.

$$N_1 + N_2 = N$$

Sample n balls at random **without replacement** ($n \leq N_1 + N_2$).

Let X be the number of red balls chosen.

Then, X is called a **hypergeometric random variable** with parameters N_1, N_2, n , and denoted by **HG**(N_1, N_2, n).

If with replacement, $\text{Bin}(n, \frac{N_1}{N_1 + N_2})$

The Hypergeometric Distribution

Example

In a small pond there are 50 fish, ten of which have been tagged.

If a fisherman's catch consists of seven fish selected at random and without replacement, and X denotes the number of tagged fish,

what is the probability that exactly two tagged fish are caught?

$$X \sim \text{HG} \left(\underset{N_1}{10}, \underset{N_2}{40}, \underset{n}{7} \right)$$

$$P(X=2) = \frac{\binom{10}{2} \binom{40}{5}}{\binom{50}{7}}$$

$$X \sim HG(N_1, N_2, n)$$

$$PMF: f(x) = \frac{\binom{N_1}{x} \binom{N_2}{n-x}}{\binom{N_1+N_2}{n}}$$

$$n-x \leq N_2$$

$$x \geq n - N_2$$

$$(x = 0, 1, 2, \dots, \min\{n, N_1\})^x$$

$$x = \max\{0, n - N_2\}, \dots, \min\{n, N_1\}$$

$$\mathbb{E} \quad N_1 = 5, \quad N_2 = 10, \quad n = 7$$

The Hypergeometric Distribution

Theorem

$$P(X = k) =$$

$$E[X] = n \frac{N_1}{N_1 + N_2}$$

$$Var[X] = n \frac{N_1}{N_1 + N_2} \frac{N_2}{N_1 + N_2} \frac{N_1 + N_2 - n}{N_1 + N_2 - 1}$$

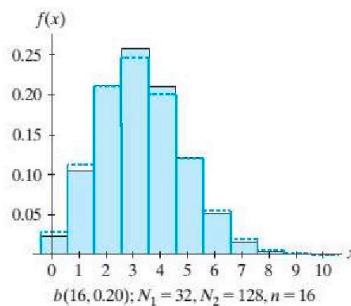
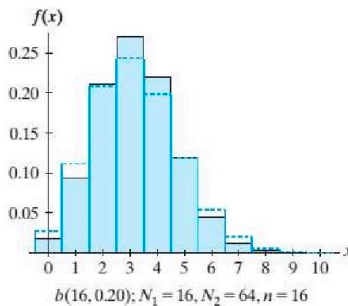
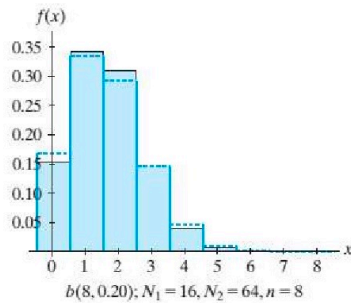
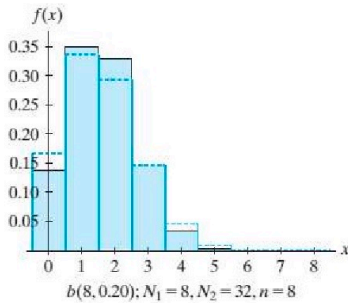
$$X \sim HG(N_1, N_2, n)$$

$$Y \sim Bin(n, \frac{N_1}{N_1 + N_2})$$

$$E[Y] = n \cdot \frac{N_1}{N_1 + N_2}$$

$$Var(Y) = n \cdot p \cdot (1-p)$$

$$= n \cdot \frac{N_1}{N_1 + N_2} \frac{N_2}{N_1 + N_2}$$



$$X \approx Y$$

when $N_1, N_2 \rightarrow \infty$

$$\frac{N_1}{N_1 + N_2} = p$$

Figure 2.5-2 Binomial and hypergeometric (shaded) probability histograms

The Hypergeometric Distribution

Exercise

In a lot (collection) of 100 light bulbs, there are five bad bulbs.

An inspector inspects ten bulbs selected at random. *without replacement*

Find the probability of finding at least one defective bulb.

$X = \#$ of Defective Bulbs chosen.

$$\sim \text{HG} \left(\frac{5}{N_1}, \frac{95}{N_2}, \frac{10}{n} \right)$$

$$P(X \geq 1) = 1 - P(X=0)$$

$$= 1 - \frac{\binom{5}{0} \binom{95}{10}}{\binom{100}{10}} \quad "$$

Section 6.

The Negative Binomial Distribution

Geometric random variables

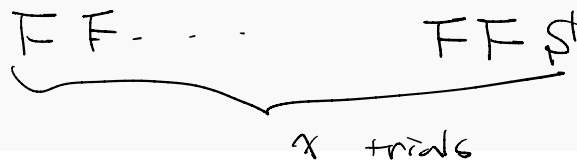
Consider a sequence of independent Bernoulli trials with success probability $p \in (0, 1)$

Let X be the number of trials **until the first success**.

This random variable is called a **geometric random variable**.

$$X \sim \text{Geom}(p)$$

$$\text{PMF : } f(x) = 1 \cdot (1-p)^{x-1} \cdot p$$
$$x = 1, 2, 3, \dots$$



Geometric random variables

$$X \sim \text{Geom}(p)$$

$$q = 1 - p$$

Theorem

The pmf of X is

$$f(x) = (1-p)^{x-1} \cdot p, \quad x = 1, 2, \dots$$

$$\mathbb{E}[X] = \frac{1}{p}$$

$$\text{Var}[X] = \frac{q}{p^2}$$

$$M(t) = \frac{pe^t}{1-(1-p)e^t}$$

Geometric random variables

Example

Some biology students were checking eye color in a large number of fruit flies.

For the individual fly, suppose that the probability of white eyes is $1/4$ and the probability of red eyes is $3/4$, and that we may treat these observations as independent Bernoulli trials.

What is the probability that at least four flies have to be checked for eye color to observe a white-eyed fly?

$X =$ # of observations until 1st white.

$\sim \text{Geom}(\frac{1}{4})$.

$$P(X \geq 4) = \sum_{x=4}^{\infty} (1-p)^{x-1} \cdot p \quad (p = \frac{1}{4}). \quad 39$$

$$= (1-p)^3 \cdot p + (1-p)^4 \cdot p + (1-p)^5 \cdot p + \dots$$

$$= \frac{(1-p)^3 \cdot p}{\underbrace{1 - (1-p)}_p} = (1-p)^3.$$

$$P(X \geq 4) = P(\text{FFF}) = (1-p)^3$$

In general, $X \sim \text{Geom}(p)$

$$P(X > k) = (1-p)^k$$

Negative Binomial random variables

$$\text{Note: } \text{Geom}(p) = \text{NegBin}(1, p)$$

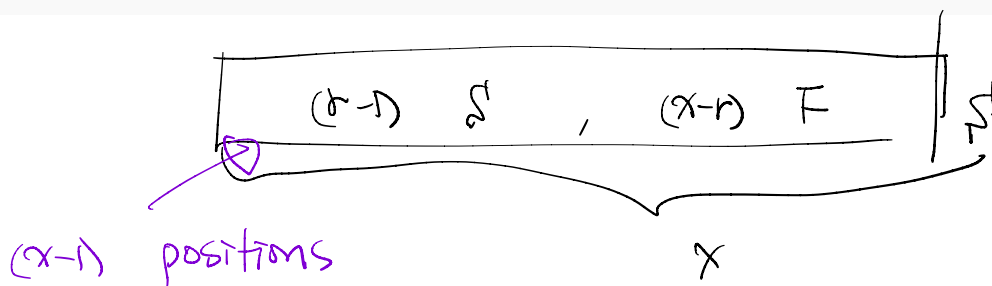
Consider a sequence of independent Bernoulli trials with success probability

Let X be the number of trials **until the r -th success**.

This random variable is called a **negative binomial random variable**.

$$X \sim \text{NegBin}(r, p)$$

$$\text{PMF: } f(x) = P(X=x) = \binom{x-1}{r-1} (1-p)^{x-r} \cdot p^r$$
$$x = r, r+1, \dots$$



Negative Binomial random variables

Theorem

The pmf of X is

$$f(k) = \binom{k-1}{r-1} p^r (1-p)^{k-r}$$

for $k = r, r+1, \dots$ and otherwise zero.

$$\mathbb{E}[X] = \frac{r}{p}$$

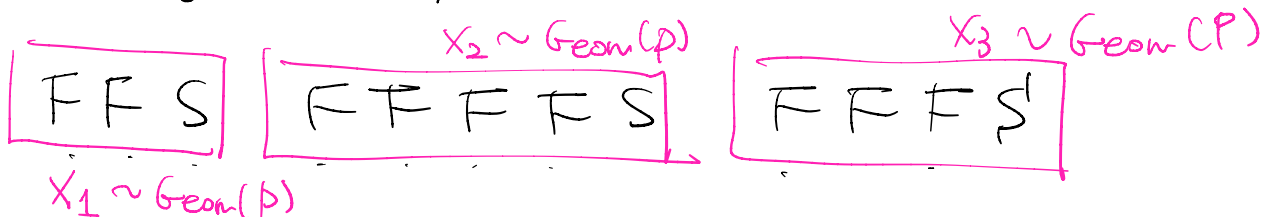
$$\text{Var}[X] = \frac{rq}{p^2}$$

$$M(t) = \left(\frac{pe^t}{1-(1-p)e^t} \right)^r$$

$$X \sim \text{NegBin}(r, p)$$

A negative binomial random variable can be written as a sum of independent geometric random variables.

Ex $X \sim \text{NegBin}(3, p)$



$$X = 12 = X_1 + X_2 + X_3$$



$$\frac{10}{(12.5)} = 80\%$$

Negative Binomial random variables

Example

Suppose that during practice a basketball player can make a free throw 80% of the time.

Furthermore, assume that a sequence of free-throw shooting can be thought of as independent Bernoulli trials.

Let X equal the minimum number of free throws that this player must attempt to make a total of ten shots.

Find the mean of X .

$$X \sim \text{Neg Bin} \left(10, \frac{4}{5} \right)$$

$$E[X] = \frac{r}{p} = 10 \cdot \frac{5}{4} = 12.5.$$

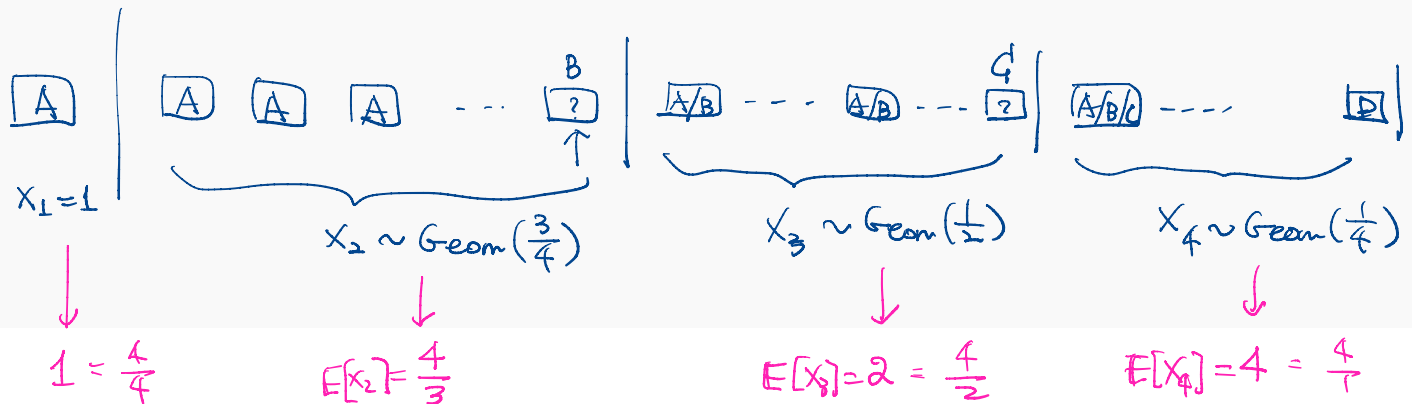
Negative Binomial random variables

Coupon Collection Problem.

Exercise

One of four different prizes was randomly put into each box of a cereal.

If a family decided to buy this cereal until it obtained at least one of each of the four different prizes, what is the expected number of boxes of cereal that must be purchased?



43

$$\begin{aligned}
 \text{ANS} &= \frac{4}{4} + \frac{4}{3} + \frac{4}{2} + \frac{4}{1} = 4 \cdot \left(\frac{1}{1} + \frac{1}{2} + \frac{1}{3} + \frac{1}{4} \right) \\
 &= \frac{25}{3}
 \end{aligned}$$

Section 7.

The Poisson Distribution

Definition

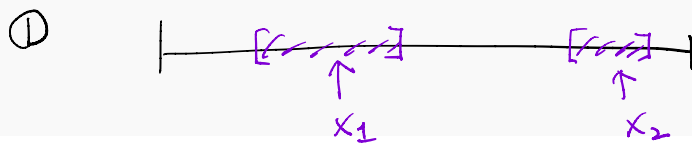
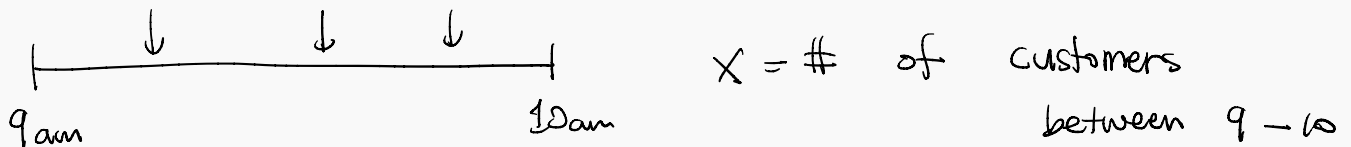
Some experiments result in counting the number of times particular events occur at given times or with given physical objects.

Example

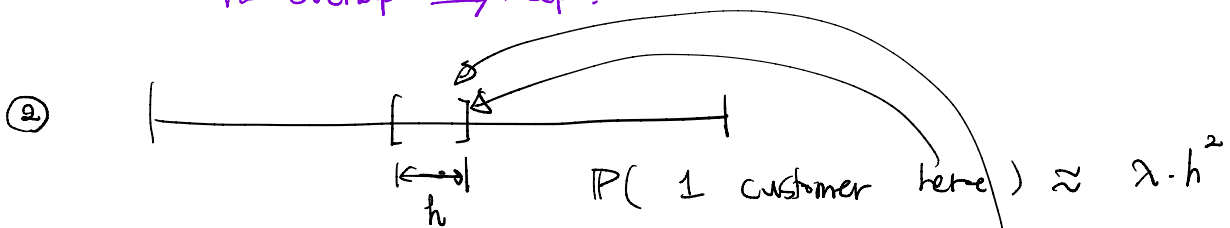
- the number of cell phone calls passing through a relay tower between 9 and 10am.
- the number of flaws in 100 feet of wire
- the number of customers that arrive at a ticket window between noon and 2pm.
- the number of defects in a 100-foot roll of aluminum screen that is 2 feet wide.

Definition

Counting such events can be looked upon as observations of a random variable associated with an **approximate Poisson process**, provided that the conditions in the following definition are satisfied.

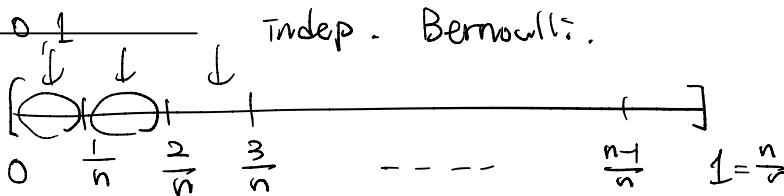


no overlap \Rightarrow indep.



③ $P(\text{more than 1 customer here}) \approx 0$

indep. Bernoulli.



$X = \# \text{ of customer} \approx \text{Bin}(n, p)$

$$E[X] = \lambda = np, \quad p = \frac{\lambda}{n}$$

$X \approx \text{Bin}(n, \frac{\lambda}{n}) \xrightarrow{n \rightarrow \infty} \text{Poisson}$

$$\binom{n}{x} \left(\frac{\lambda}{n}\right)^x \left(1 - \frac{\lambda}{n}\right)^{n-x} \rightarrow e^{-\lambda} \frac{\lambda^x}{x!}$$

Definition

Let the number of occurrences of some event in a given continuous interval be counted. Then we have an **approximate Poisson process** with parameter $\lambda > 0$ if

- The numbers of occurrences in nonoverlapping subintervals are independent.
- The probability of exactly one occurrence in a sufficiently short subinterval of length h is approximately λh .
- The probability of two or more occurrences in a sufficiently short subinterval is essentially zero.

Under these assumption, consider the number of occurrences in a time interval $[0, 1]$.

Definition

Split $[0, 1]$ into n subintervals $[0, \frac{1}{n}]$, $[\frac{1}{n}, \frac{2}{n}]$, \dots , $[\frac{n-1}{n}, 1]$.

In each subinterval, at most one event occurs with probability $\frac{\lambda}{n}$.

Thus, the number of occurrences is a binomial random variable with n ~~and~~ ^{and} $\frac{\lambda}{n}$.

As $n \rightarrow \infty$, the random variable gets close to some random variable X .

We say X is a **Poisson random variable with parameter λ** if its pmf is

$$\mathbb{P}(X = k) = \frac{e^{-\lambda} \lambda^k}{k!}$$

for $k = 0, 1, 2, \dots$.

Definition

$$X \sim \text{Pois}(\lambda)$$

$$f(k) = e^{-\lambda} \frac{\lambda^k}{k!}$$

$k = 0, 1, 2, \dots$

Theorem

$$E[X] = \lambda$$

$$\text{Var}[X] = \lambda$$

$$M(t) = e^{\lambda(e^t - 1)}$$

$$\begin{aligned}
 e^x &= \sum_{n=0}^{\infty} \frac{x^n}{n!} \\
 E[X] &= \sum_{k=1}^{\infty} k \cdot e^{-\lambda} \frac{\lambda^k}{k!} = e^{-\lambda} \cdot \sum_{k=1}^{\infty} \frac{\lambda^{(k-1)} \cdot \lambda}{(k-1)!} \\
 &= \lambda \cdot e^{-\lambda} \left(\sum_{n=0}^{\infty} \frac{\lambda^n}{n!} \right) = \lambda \underbrace{e^{-\lambda}} \underbrace{e^{\lambda}} = \lambda.
 \end{aligned}$$

$$X = \# \text{ of customers between } 9-10. \sim \text{Pois}(\lambda)$$

$$\cdot f(k) = e^{-\lambda} \frac{\lambda^k}{k!} \quad k = 0, 1, 2, \dots$$

$$\cdot E[X] = \lambda, \quad \text{Var}(X) = \lambda, \quad M(t) = e^{\lambda(e^t - 1)}$$

Definition

Example

In a large city, telephone calls to 911 come on the average of two every 3 minutes.

If one assumes an approximate Poisson distribution, what is the probability of five or more calls arriving in a 9 minute period?

$$X = \# \text{ of calls in } 9 \text{ min} \sim \text{Pois}(6)$$

$$P(X \geq 5) = \sum_{k=5}^{\infty} f(k) = \sum_{k=5}^{\infty} e^{-6} \cdot \frac{6^k}{k!}$$

↑
use table

Poisson Approximation to Binomial

Suppose X is a binomial random variable $b(n, p)$, n is large, and p is small but np converges to some constant, say λ .

In this case, X can be approximated by a Poisson random variable with parameter λ .

This approximation is quite accurate if $n \geq 20, p \leq 0.05$ or $n \geq 100, p \leq 0.1$.

$$X \sim \text{Bin}(n, p) \approx \text{Pois}(\lambda)$$

$\uparrow \quad \uparrow$
large small

$\lambda = np$

Ex

$$X \sim \text{Bin}(1000, 0.99) \quad X$$
$$1000 - X = Y \sim \text{Bin}(1000, 0.01) \approx \text{Pois}(10)$$

Poisson Approximation to Binomial

Example

A manufacturer of Christmas tree light bulbs knows that 2% of its bulbs are defective. Assuming independence, ^Xthe number of defective bulbs in a box of 100 bulbs has a binomial distribution with parameters $n = 100$ and $p = 0.02$.

Find the probability that a box of 100 of these bulbs contains at most three defective bulbs.

$$\begin{aligned} X &\sim \text{Bin}(100, 0.02) \\ P(X \leq 3) &= \sum_{k=0}^3 \binom{100}{k} (0.02)^k (0.98)^{100-k} \\ &\approx \sum_{k=0}^3 e^{-2} \frac{2^k}{k!} \\ &= e^{-2} \cdot \left(\frac{2^0}{0!} + \frac{2^1}{1!} + \frac{2^2}{2!} + \frac{2^3}{3!} \right) \\ &= e^{-2} \cdot \left(1 + 2 + 2 + \frac{8}{6} \right) = \frac{14}{3} e^{-2} \\ &\approx 0.857. \\ &\uparrow \\ &\text{table.} \end{aligned}$$

Poisson Approximation to Binomial

Exercise

Suppose that the probability of suffering a side effect from a certain flu vaccine is 0.005. If 1000 persons are vaccinated, approximate the probability that (a) At most one person suffers. (b) Four, five, or six persons suffer.

Ch1 - #10, 5, 4

S #5.

#10 2R, 4W Sample 5 with replacement.

P(No two balls drawn consecutively have the same color)

$$= P(WRWRW \text{ or } \underline{RWRWR})$$

$$= 1 \cdot \left(\frac{2}{6}\right)^2 \cdot \left(\frac{4}{6}\right)^3 + 1 \cdot \left(\frac{2}{6}\right)^3 \cdot \left(\frac{4}{6}\right)^2$$

$$= \left(\frac{1}{3}\right)^2 \cdot \left(\frac{2}{3}\right)^3 + \left(\frac{1}{3}\right)^3 \cdot \left(\frac{2}{3}\right)^2 = \frac{8+4}{3^5} = \frac{12}{243}$$

without replacement

$$= P(WRWRW) = \frac{4}{6} \cdot \frac{2}{5} \cdot \frac{3}{4} \cdot \frac{1}{3} \cdot \frac{2}{2} = \frac{1}{15}$$

#5 5G, 7R → Sample 9.

A = { First 3 balls = G }

B = { Exactly 4 G } = { 4G, 5R }

(a) With Replacement: P(B), P(A ∩ B) = ?

$$P(B) = P(X=4) = \binom{9}{4} \left(\frac{5}{12}\right)^4 \cdot \left(\frac{7}{12}\right)^5$$

X = # of G in 9 sample ~ Bin(9, 5/12)

Note: P(A) = $\left(\frac{5}{12}\right)^3$

P(A ∩ B) = P(A) · P(B) ? ← No. A, B Not indep.

$$P(A \cap B) = \binom{6}{1} \left(\frac{5}{12}\right)^4 \cdot \left(\frac{7}{12}\right)^5$$

GGG 6
↑
1G, 5R

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{\binom{6}{1}}{\binom{9}{4}} = \frac{6}{\frac{9 \cdot 8 \cdot 7 \cdot 6}{4!}} = \frac{1}{21}$$

(b) without replacement.

$$P(B) = P(Y=4) = \frac{\binom{5}{4} \binom{7}{5}}{\binom{12}{9}}$$

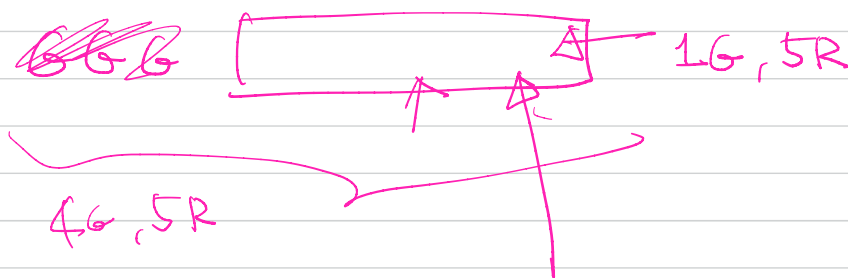
↑
4G, 5R

Y = # of G in 9 samples

~ HG(5, 7, 9)

$$P(A) = \frac{5}{12} \cdot \frac{4}{11} \cdot \frac{3}{10} = \frac{1}{22}$$

$$P(B|A) = P(Z=1) = \frac{\binom{2}{1} \binom{7}{5}}{\binom{9}{6}}$$



Z = # of G in last 6 ~ HG(2, 7, 6)

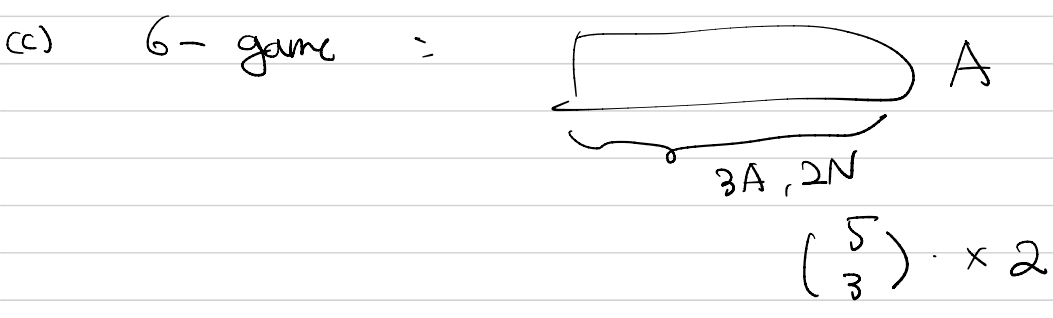
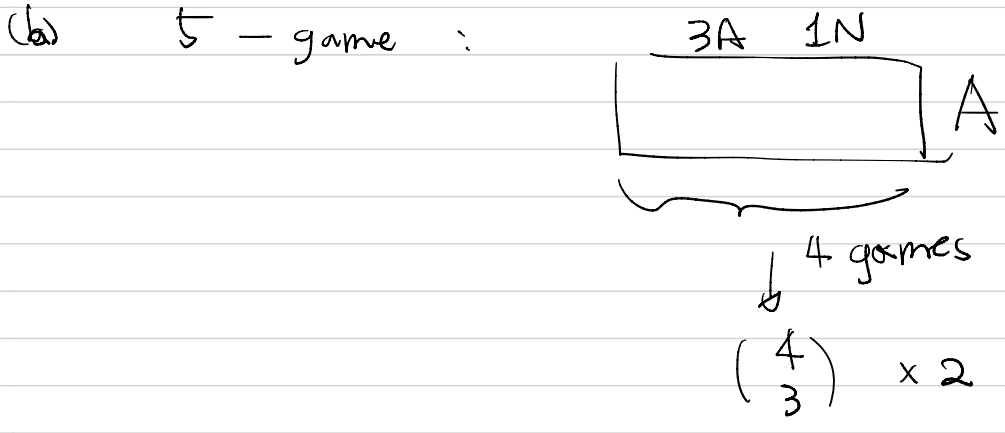
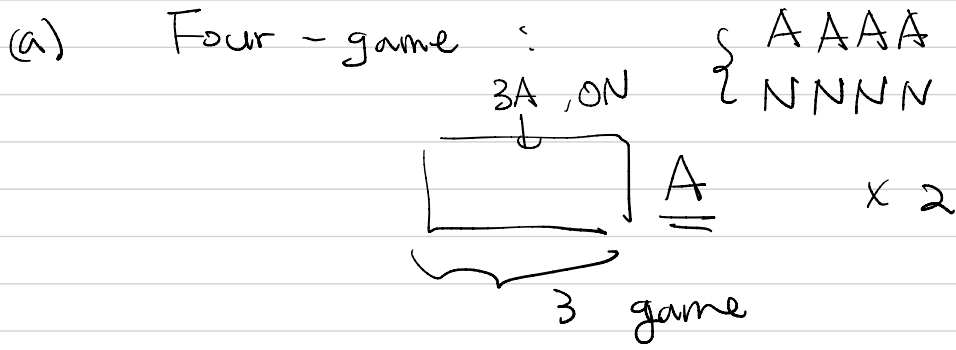
#5

$$P(A|B) = \frac{P(B|A) \cdot P(A)}{P(B)}$$

#4

WS in baseball A, N

win 4 games first → winner.



Sample #5

$$M(t) = \underline{c(1-2t)^{-5}}, \quad t < \frac{1}{2}$$

(a) $c = ?$

$$M(t) = \mathbb{E}[e^{tX}]$$

$$M(0) = \mathbb{E}[e^0] = 1 = c(1-0)^{-5} = c \quad \therefore c=1.$$

(b) $\mathbb{E}[X] = M'(0) = (-5) \cdot (1-2t)^{-6} \cdot (1-2t)' \Big|_{t=0}$
 $= (-5) \cdot (-2) \cdot \underbrace{(1-2t)^6}_{=1} \Big|_{t=0}$
 $= 10.$

(c) $\text{Var}(X) = \mathbb{E}[X^2] - (\mathbb{E}[X])^2 = 20 - (10)^2 = 20.$

$$\mathbb{E}[X^2] = M''(0) = 10 \cdot (-6) \cdot (1-2t)^{-7} \cdot (-2) \Big|_{t=0}$$
$$= 120$$

Sample #4

$$f(k) = \begin{cases} c & k=0 \\ \frac{1}{3^k k!} & k=1, 2, \dots \end{cases}$$

$c = ?$

$$1 = \sum_{k=0}^{\infty} f(k) = c + \underbrace{\sum_{k=1}^{\infty} \frac{1}{k!} \left(\frac{1}{3}\right)^k}_{e^{\frac{1}{3}} - 1}$$
$$2 - e^{\frac{1}{3}} = c.$$

$$e^x = \sum_{k=0}^{\infty} \frac{x^k}{k!}$$

$$(a+b)^n = \sum_{k=0}^n \binom{n}{k} a^k b^{n-k}$$

→ Binomial

$$a + a \cdot r + ar^2 + \dots$$

→ Geom.

$$e^x = \sum_{k=0}^{\infty} \frac{x^k}{k!}$$

→ Poisson.